

DRAFT
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Advanced Statistical Analysis of Multi-location Variety Trials

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Statistical Analysis of Multi- location Variety Trials

This Module

Objectives and Output

To develop the participant 's skill in designing multi-locational plant variety trials, carrying out the statistical analyses of the data generated from such designs, interpretation and presentation of the results from such analyses. Each participant will be expected to a prepare a draft scientific manuscript using the data from his/her own experiments.

Scope

The course covers:

- Design and analysis of data from Multi-locational Variety Trials and interpretation of GxE interaction, in general. In particular it focuses on - analysis of data from individual environments, test for homogeneity of error variances, combined analysis of data (evaluation of GxE interaction and tests for parallelism of the regression lines), common stability statistics, clustering methods (hierarchical and non-hierarchical cluster analysis of genotypes/environments), principal component analysis (of genotypes/environments) and heritability of the traits in broad sense (from individual environments as well as all the environments combined), additive main-effects and multiplicative interaction model, and inter-site transferability of crop varieties.
- Programs coded in GENSTAT 5 Release 4.2 for analysing data from multi-locational trials conducted in block designs.

In this note we have adapted the materials from several sources and have been cited in the reference/bibliography.

Statistical Analysis of Multi- locational Variety Trials

1. Introduction

Crop improvement process is long and involved several stages of germplasm collections, selection of desired material types, developing crosses between desired parents, preliminary evaluation of (generally large number of) genotypes/lines, selection, evaluations of selected genotypes in replicated trials followed by further (advanced) yield trials at multi-environments/locations representing the target domain for which the varieties are ultimately developed for production. When a number of varieties of a crop is grown over several environments (locations, years), their relative responses on various characters may show variation over the environments. This happens due to differential interplay between genetic and non-genetic (environmental) factors and therefore such a variation is said to occur due to interaction between genotype and environment or genotype x environment interaction (GEI). The major aspect of multi-locational trials is to identify stable and/or adaptable genotypes to the changing environments. Search for such genotypes requires a careful examination and exploitation of the GEI.

Keeping above aspect in view, this manuscript discusses commonly used experimental designs, data analysis from individual environments, combined analysis of data in Section 2, stability analyses Section 3, partitioning of GEI in Section 4, stochastic dominance in Section 5, a brief introduction to additive main-effects and multi-plicative interaction (Section 6) and inter-site transferability of crop varieties (Section 7). Various analyses covered in these sections have been coded in GENSTAT 5 (Genstat 5 Committee 1993) and a sample printout has been presented in Section 8.

2. Estimation of genotype means and GEI

2.1 Experimental design and data

We consider for that a set of p genotypes have been evaluated in q environments in replicated trials conducted in complete or incomplete block designs. A checklist of concerns in planning of an experiment is given by Jeffers (1978). The number of replications at a site depends on the variability in the experimental material (plots) and on the precision required of the estimates. For multi-locational trials, Kempthorne (1952, p583) provides an expression for an optimum number of replication in terms of error

variance, genotype variance and genotype x environment interaction variance and the cost factors. Although two replicates are absolutely minimum to estimate experimental error variance, in many case this is also the optimum number.

A number of statistical packages, such as GENSTAT 5, ALPHAGEN, ALPHA+, GENDEX, etc. can facilitate generating randomized plans for various types of experimental designs. Plot-wise records on the response variable, generally taken as yield, are required for statistical analysis.

2.2. Analysis of variance from individual locations

Analysis of variance would be generated by fitting the model

yield = general mean + genotype effect + replicate (or complete block) effect + error

for randomized complete block design (RCBD), and

yield = general mean + genotype effect + replicate effect
+ effect of incomplete block within replicate + error

for the incomplete block design used.

The above model for data from RCBDs can be expressed using the following notations.

$$y_{ij} = \mu + \tau_i + b_j + e_{ij}$$

where y_{ij} = yield corresponding to

j – th block ($j = 1, 2, \dots, r$), i – th variety ($i = 1, 2, \dots, v$);

μ = general mean, τ_i = effect of i – th genotype, b_j = effect of j – th block,

e_{ij} = normally and independently distributed random variables with mean zero

and

variance σ^2 .

If the varieties are selected randomly to represent a population, we then assume that the τ_i s are normally and independently distributed random variables with mean zero and variance σ_g^2 . Analysis of variance for the data with expected mean squares is in the following:

Table. Analysis of variance and expected mean squares from a single location data

Source	df	Mean Square	Expectation of Mean Square
<i>(a) Variety effects assumed fixed.</i>			
Blocks	$r - 1$	—	—
Varieties	$v - 1$	V	$\sigma^2 + \frac{r}{v-1} \sum (\tau_i - \bar{\tau})^2$
Error	$(r - 1)(v - 1)$	E	σ^2
<i>(b) Variety effects assumed random</i>			
Blocks	$r - 1$	—	—
Varieties	$v - 1$	V	$\sigma^2 + r \sigma_g^2$
Error	$(r - 1)(v - 1)$	E	σ^2

The estimates of variance components are

$$\hat{\sigma}^2 = E$$

$$\hat{\sigma}_g^2 = (V - E)/r$$

Gain due to selection

The gain in selecting a chosen proportion p of the lines is the difference between mean of the selected lines and the population mean. This difference has the expected value

$$= \frac{K \sigma_g^2}{\sqrt{\sigma^2 + \sigma_g^2}}$$

The constant K , a function of p , is given by

$K = Z/p$; Z = ordinate of the standard normal distribution at the point which covers the area p in the right tail of the distribution and is given by

$$Z = (1/(2\pi)^{1/2}) \exp(-x^2/2)$$

where $p = \int_x^\infty (1/(2\pi)^{1/2}) \exp(-u^2/2) du$.

For example $p = 0.2$ (i.e. 20%) gives $Z = 10.28$ and $K = 1.40$.

If the selection is based on means of r replications then the expected difference is

$$= \frac{K \sigma_g^2}{\sqrt{\sigma_g^2 + \sigma^2/r}}$$

Models can similarly be written for other type of experimental designs. These models can be fitted using statistical packages (e.g. GENSTAT, SAS). The analysis of variance provides an estimate of experimental errors variance and a test of significance if the varietal differences in the yield response are real rather than arising from experimental error or chance. Means of varieties can be estimated (with adjustment for incomplete blocks where used) for individual environments. The experimental error variances may be examined for their homogeneity over the environments using Bartlett's chi-square test. If the error variances are found homogeneous then a pooled error variance can be obtained.

2.3 Combined analysis of variance over all locations

The combined analysis of variance to study GEI should distinguish the two cases- i) error variances homogeneous and ii) error variances heterogeneous.

Homogeneous error variances

One may estimate the GEI by fitting the model for RCBDs

Yield = General mean + Environment effect + Replication effects (within environments) + genotype effect + genotype x environment interaction effect + error

This will produce a common (pooled) error for testing significance of GEI.

Expressing the above model in notations

$$y_{ijk} = \mu + E_j + \beta_{kj} + \tau_i + (GE)_{ij} + e_{ijk} \quad (8)$$

where

variety, $i = 1, 2, \dots, v$

environment $j = 1, 2, \dots, L$

blocks $k = 1, 2, \dots, r$

E_j = effect of j - th environment; β_{kj} = effect of k - th block in j - th

environment;

τ_i = effect of i - th variety;

$(GE)_{ij}$ = interaction term for i - th genotype and j - th environment;

e_{ijk} = plot error assumed independent and normally distributed with mean zero

and

constant variance σ^2 .

Let us assume that the environment and variety effects are fixed and the genotype x environment interaction are independent and normally distributed with means zero and variance σ_{ge}^2 . We have the following ANOVA structure.

Table. Analysis of variance of data from several experiments conducted in RCBDs

<i>Source Square</i>	<i>df</i>	<i>Mean Square</i>	<i>Expectation of Mean</i>
Environment	$L - 1$	—	—
Blocks within Envs.	$L(r - 1)$	—	—
Variety	$v - 1$	V	
$\sigma^2 + r\sigma_{ge}^2 + \frac{rL}{v-1} \sum (\tau_i - \bar{\tau})^2$			
Variety \times Envs.	$(v - 1)(L - 1)$	I	$\sigma^2 + r\sigma_{ge}^2$
Error	$L(r - 1)(v - 1)$	E	σ^2
Total	$Lrv - 1$		

(Assuming environment effects fixed, genotype effects and the interaction effects random)

Envs. =Environments

Gain due to selection

The gain in selecting a chosen proportion p of the lines (based on means over the replications) has the expected value

$$= \frac{K\sigma_g^2}{\sqrt{\sigma_g^2 + \sigma_{ge}^2/L + \sigma^2/(Lr)}}$$

Example 1. Combined ANOVA over RCBDs

Data: Chickpea yields

Identifier	Type	Length	Values	Missing
Location	Factor	460	Present	0
Rep	Factor	460	Present	0
Geno	Factor	460	Present	0
Yield	Variate	460	Present	0

***** Analysis of variance *****

Variate: Yield

Source of variation	d.f.	s.s.	m.s.	v.r.	F pr.
Location.Rep stratum					
Location	4	1.127E+08	2.817E+07	22.10	<.001
Residual	15	1.911E+07	1.274E+06	8.30	
Location.Rep.Geno stratum					
Geno	22	7.644E+06	3.474E+05	2.26	0.001
Location.Geno	88	3.968E+07	4.509E+05	2.94	<.001
Residual	330	5.064E+07	1.535E+05		
Total	459	2.297E+08			

***** Tables of means *****

Variate: Yield

Grand mean 1016.

Location	1.00	2.00	3.00	4.00	5.00		
	1012.	322.	880.	1867.	998.		
Geno	1.00	2.00	3.00	4.00	5.00	6.00	7.00
	1124.	976.	942.	854.	1057.	1254.	979.
Geno	8.00	9.00	10.00	11.00	12.00	13.00	14.00
	1128.	951.	959.	1196.	1121.	733.	1040.
Geno	15.00	16.00	17.00	18.00	19.00	20.00	21.00
	1073.	1081.	856.	1030.	1028.	1186.	1069.
Geno	22.00	23.00					
	761.	965.					
Location	Geno	1.00	2.00	3.00	4.00	5.00	6.00
1.00		828.	1054.	1016.	1023.	880.	1111.
2.00		172.	284.	401.	132.	381.	312.
3.00		933.	792.	750.	625.	818.	917.
4.00		2579.	1954.	1852.	1238.	2207.	2931.
5.00		1107.	798.	691.	1250.	1000.	1000.
Location	Geno	7.00	8.00	9.00	10.00	11.00	12.00
1.00		773.	1078.	918.	1240.	1153.	1000.
2.00		328.	555.	281.	276.	318.	427.
3.00		932.	875.	766.	826.	792.	719.
4.00		1875.	2598.	2149.	1704.	2942.	2556.
5.00		988.	536.	643.	750.	774.	905.
Location	Geno	13.00	14.00	15.00	16.00	17.00	18.00
1.00		979.	905.	1005.	943.	882.	1036.
2.00		271.	293.	326.	481.	286.	349.
3.00		818.	693.	906.	875.	891.	1110.
4.00		906.	2285.	2117.	1783.	1446.	1729.
5.00		691.	1024.	1012.	1322.	774.	929.
Location	Geno	19.00	20.00	21.00	22.00	23.00	
1.00		1183.	1262.	1059.	926.	1014.	
2.00		280.	333.	354.	228.	339.	
3.00		1099.	1219.	1026.	959.	896.	
4.00		803.	1903.	1442.	929.	1016.	
5.00		1774.	1214.	1465.	762.	1560.	

*** Standard errors of means ***

Table	Location	Geno	Location Geno
rep.	92	20	4
e.s.e.	117.7	87.6	224.8
d.f.	15	330	151.47
Except when comparing means with the same level(s) of			
Location			195.9

d.f. 330

*** Standard errors of differences of means ***

Table	Location	Geno	Location Geno
rep.	92	20	4
s.e.d.	166.4	123.9	318.0
d.f.	15	330	151.47

Except when comparing means with the same level(s) of
Location 277.0
d.f. 330

*** Least significant differences of means ***

Table	Location	Geno	Location Geno
rep.	92	20	4
l.s.d.	354.7	243.7	628.2
d.f.	15	330	151.47

Except when comparing means with the same level(s) of
Location 544.9
d.f. 330

Example 2. Combined analysis of data from triple lattices

Data: *Barley yield*: $v = 64$ genotypes, 10 environments.

***** REML Variance Components Analysis *****

Response Variate : Yield

Fixed model : Constant+Loc+Geno+Loc.Geno
Random model : Loc.Rep+Loc.Rep.Blk

Number of units : 1920
No absorbing factor

*** Estimated Variance Components ***

Random term	Component	S.e.
Loc.Rep	34618.	13020.
Loc.Rep.Blk	41528.	5621.
units	83456.	3642.

*** Approximate stratum variances ***

		Effective d.f.
Loc.Rep	2631205.	20.00
Loc.Rep.Blk	304939.	210.00
units	83456.	1050.00

* Matrix of coefficients of components for each stratum *

Loc.Rep	64.00	8.00	1.00
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Loc.Rep.Blk	0.00	5.33	1.00
units	0.00	0.00	1.00

*** Wald tests for fixed effects ***

Fixed term	Wald statistic	d.f.
Loc	1521.7	9
Geno	517.4	63
Loc.Geno	1314.0	567

*** Table of predicted means for Constant ***

*** Table of predicted means for Geno ***

Geno	1	2	3	4	5
	2132	2332	2251	2315	2028
Geno	6	7	8	9	10
	1955	1993	2228	2209	1936
Geno	11	12	13	14	15
	2099	2314	2261	2288	2093
Geno	16	17	18	19	20
	2079	1969	2102	2124	2131
Geno	21	22	23	24	25
	2108	2113	1663	1916	1784
Geno	26	27	28	29	30
	1980	2166	1738	2027	1951
Geno	31	32	33	34	35
	2061	2333	1933	2296	1861
Geno	36	37	38	39	40
	2167	1946	2273	2247	2314
Geno	41	42	43	44	45
	2243	2189	2068	2164	2216
Geno	46	47	48	49	50
	2146	2159	2302	2327	2154
Geno	51	52	53	54	55
	2283	2296	2260	1847	2257
Geno	56	57	58	59	60
	2304	2335	2245	2300	2263
Geno	61	62	63	64	
	2288	2164	2194	2176	

Standard error of differences:	Average	78.98
	Maximum	79.51
	Minimum	78.23

Average variance of differences: 6237.

*** Table of predicted means for Loc.Geno ***

Geno	1	2	3	4	5
Loc					
1	421	360	379	357	353
2	819	715	716	759	808
3	1980	2442	2372	1927	2305
4	1180	1379	1090	996	1225
5	3327	3716	3386	3858	2737
6	4422	4764	4100	4935	3178
7	1080	1142	1423	1246	1302
8	4918	5038	5160	4872	5052
9	1570	1485	1712	2191	1296
10	1601	2278	2171	2008	2027
.....					
Geno	61	62	63	64	
Loc					
1	430	533	416	475	
2	805	989	823	864	
3	2325	2501	2374	2237	
4	1246	1102	1196	778	
5	3347	3046	3236	3774	
6	4167	3646	4181	4216	

7	1324	825	1264	1305
8	5102	5413	5013	4410
9	1701	1437	1412	1599
10	2436	2149	2024	2100
Standard error of differences:		Average	293.2	
		Maximum	298.0	
		Minimum	246.4	
Average variance of differences:			86200.	
Standard error of differences for same level of factor:				
	Loc	Geno		
Average	249.7	298.0		
Maximum	251.4	298.0		
Minimum	246.4	298.0		
Average variance of differences:				
	62372.	88806.		

2.4 Combined analysis of data from experiments in RCBDs conducted over several locations and years

We shall consider the analysis of a trial on v varieties evaluated in r randomized blocks at each of L locations in each of the same Y years. We shall use is the following model.

$$y_{ijkl} = \mu + L_j + Y_k + (LY)_{jk} + \tau_i + (L\tau)_{ij} + (Y\tau)_{ik} + (LY\tau)_{ijk} + \beta_{jkl} + e_{ijkl}$$

where the various terms in right hand side represent general mean, location effect, year effect, location x year interaction, variety effects, location x variety interaction, year x variety interaction, location x year x variety interaction, blocks within location and year, and plot error respectively and associated with the suffixes representing the following:

$i = 1, 2, \dots, v$: variety

$j = 1, 2, \dots, L$: locations

$k = 1, 2, \dots, T$: years

$l = 1, 2, \dots, r$: blocks.

Further we assume that variety and replication effects are fixed while location effect and all other factor effects are random. The interactions with variety i.e. location x variety interaction, year x variety interaction, location x year x variety interaction, and the plot errors are assumed independently and normally distributed with means zero and variances σ_{lv}^2 , σ_{yg}^2 , σ_{lyg}^2 and σ^2 respectively. We have the following ANOVA structure.

Table Analysis of variance of experiments combined over several locations and years

<i>Source</i>	<i>df</i>	<i>Mean Square</i>	<i>Expectation of Mean Square</i>
Location	$L - 1$	Lo	—
Years	$Y - 1$	Ye	—
Location × years	$(L - 1)(Y - 1)$	LoYe	—
Varieties	$v - 1$	V	$\sigma^2 + r\sigma_{lyg}^2 + rL\sigma_{yg}^2 + rY\sigma_{ig}^2 + \frac{rLY}{v-1} \sum (\tau_i - \bar{\tau})^2$
Places × varieties	$(L - 1)(v - 1)$	LoV	$\sigma^2 + r\sigma_{lyg}^2 + rY\sigma_{ig}^2$
Years × varieties	$(Y - 1)(v - 1)$	YeV	$\sigma^2 + r\sigma_{lyg}^2 + rL\sigma_{yg}^2$
Places × years × varieties	$(L - 1)(Y - 1)(v - 1)$	LoYeV	$\sigma^2 + r\sigma_{lyg}^2$
Replications	$LY(r - 1)$	—	
Error	$LY(v - 1)(r - 1)$	E	σ^2
Total	$LYvr - 1$		

Gain due to selection

The gain in selecting a chosen proportion p of the lines (based on means over the replications) has the expected value

$$= \frac{K\sigma_g^2}{\sqrt{\sigma_g^2 + \sigma_{ig}^2/L + \sigma_{yg}^2/Y + \sigma_{lyg}^2/(LY) + \sigma^2/(LYr)}}$$

Heterogeneous error variances

We can fit the following model on the genotype x environment data on means (or adjusted means for incomplete blocks) using a weighted analysis of variance with weights being inversely proportional to the variance of the means. The weight corresponding to a mean (for a combination of genotype and environment) may be estimated by $1/(\text{standard error of the mean})^2$.

$$\text{Mean} = \text{Environment effect} + \text{Genotype effect} + \text{residual}$$

The residual sum of squares produced by the weighted least-squares would be the weighted GEI sum of squares and would be approximately distributed as chi-square with GEI degrees of freedom.

Once there is a significant GEI, we may carry out further analyses to identify the causes of interaction inters of the responsiveness of the genotypes to the environments.

3 Exploitation of G x E Data: Stability Analysis

Genotypes performance changes due to environmental pressures or stresses (due to the population heterogeneity or population buffering and changes in the genetic make up taking place over generations) and differences in their ability to adapt to the stress factors (short-term acclimatization). A number of statistical models to study genotypic adaptation based on phenotypic performance have been discussed in literature. Byth and Mungomery (1981) discussed the following three concepts *Stability, adaptability, and predictability*.

"

Phenotypic *stability* refers to the ability of a genotype to maintain a near constant phenotype for the character of interest over variable environments. Such a genotype would be regarded as having *wide adaptation*. But certain genotypes may also show predictably superior performance in particular types of environments indicating that broad adaptation inevitably involves sacrifice of performance in specific environments. Thus the strategies of plant improvements for broad adaptation (minimizing G x E interaction) and specific adaptation (emphasizing favorable interaction) are in direct conflict.

Predictability refers to the extent to which response is systematic.

Responsiveness is the ability of a genotype to respond in a particular manner to a general change in the environmental potential.

Sensitivity (also *stability*) refers to the extent of unpredictable variation in response. Some researchers relate stability to variability of performance over time (temporal variation) at a location while adaptability to variability in performance across locations (spatial variation).

"

We shall in the present chapter discuss various concepts of stability using statistical measures in common practice. The various concepts and measures of stability, originating due to different outlooks of experimenter to their specific problems, have added to the difficulty of choosing a stability parameter (s) for a given situation. We include the two approaches discussed by Lin et al (1986).

3.1 Parametric Approach

Stability statistics are derived (computed) for each genotype from two-way tables of genotype and environment data. These statistics are based on either of the following three types of stability concepts. A genotype is considered to be stable if

- its among environments variance is small (Type I stability),
- its response environments is parallel to mean response of all genotypes in trial (Type II stability),
- the residual mean square from regression model on environment index is small (Type III stability).

In order to list various statistics, we shall use the following notations. Let y_{ij} denote the mean value of i -th genotype in the j -th environment ($i=1,2, \dots, p$, $j=1,2,\dots,q$). Let

$$\bar{y}_{i.} = \sum_j y_{ij}/q; \quad \bar{y}_{.j} = \sum_i y_{ij}/p, \quad \bar{y}_{..} = \sum_i \sum_j y_{ij}/(pq)$$

represent respectively, means of i -th genotype, j -th environment and overall mean. The nine statistics and one more in current use are briefly described as follows:

1. The variance of a genotype across environments

$$S_i^2 = \sum_{j=1}^q (y_{ij} - \bar{y}_{i.})^2 / (q-1),$$

2. The coefficient of variation

$$CV_i = S_i / \bar{y}_{i.}$$

Francis and Kannenberg (1978) used the conventional CV% of each genotype as a stability measure.

3. Plaisted and Peterson's (1959) mean variance component for pair-wise G x E interaction ($\bar{\theta}_i$)

$$\bar{\theta}_i = (p \sum_{j=1}^q (y_{ij} - \bar{y}_{i.} - \bar{y}_{.j} + \bar{y}_{..})^2 + \sum \sum (y_{ij} - \bar{y}_{i.} - \bar{y}_{.j} + \bar{y}_{..})^2) / (2(p-1)(q-1)).$$

The mean of the estimated variance components of the G x E interaction for all pairs of genotypes that include genotype i is the stability measure of genotype i.

4. Plaisted's (1960) variance component for G x E interaction ($\theta_{(i)}$)

$$\theta_{(i)} = (-p \sum_{j=1}^q (y_{ij} - \bar{y}_{i.} - \bar{y}_{.j} + \bar{y}_{..})^2 / (p-1) + \sum \sum (y_{ij} - \bar{y}_{i.} - \bar{y}_{.j} + \bar{y}_{..})^2 / ((p-2)(q-1))).$$

One genotype i is deleted from the entire set of data and the G x E interaction variance from this subset is the stability index for genotype i.

5. Wricke's (1962) ecovalence (w_i^2)

$$w_i^2 = \sum_{j=1}^q (y_{ij} - \bar{y}_{i.} - \bar{y}_{.j} + \bar{y}_{..})^2.$$

This G x E interaction effects for genotype i, squared and summed across all environments, is the stability measure for genotype i.

6. Shukla's (1972a) stability variance (σ_i^2)

$$\sigma_i^2 = (p \sum (y_{ij} - \bar{y}_{i.} - \bar{y}_{.j} + \bar{y}_{..})^2 - \sum \sum (y_{ij} - \bar{y}_{i.} - \bar{y}_{.j} + \bar{y}_{..})^2 / (p-1)) / ((p-2)(q-1)).$$

Based on residuals in a two-way classification, the variance of a genotype across environments is the stability measure.

7. Finlay and Wilkinson's (1963) regression coefficients (b_i)

$$b_i = \sum_j (y_{ij} - \bar{y}_{i.})(\bar{y}_{.j} - \bar{y}_{..}) / \sum_j (\bar{y}_{.j} - \bar{y}_{..})^2.$$

The observed values are regressed on environmental indices environments and the overall means. The regression coefficient of each genotype is taken as its stability parameter.

8. Perkins and Jinks' (1968) regression coefficient (β_i)

$$\beta_i = \sum_j (y_{ij} - \bar{y}_{i.} - \bar{y}_{.j} + \bar{y}_{..})(\bar{y}_{.j} - \bar{y}_{..}) / \sum_j (\bar{y}_{.j} - \bar{y}_{..})^2.$$

Similar to (7) except that the observed values are adjusted for environment effects for computing regression coefficients ($\beta_i = b_i - 1$).

9. Eberhart and Russell's (1966) deviation parameter (δ_i^2)

$$\delta_i^2 = (\sum_j (y_{ij} - \bar{y}_i - \bar{y}_{.j} + \bar{y}_{..})^2 - \beta_i^2 \sum_j (\bar{y}_{.j} - \bar{y}_{..})^2) / (q-2)$$

This is the residual mean square (MS) of deviation from regression defined in (7) or (8) is the measure of stability.

10. Variance of genotypes across environments on the ratios of yields to environment means

Yau (1972) gave an other statistics, denoted here by ξ_i as the variance across environments of the ratios of yields to the mean under respective environment

$$\xi_i = \sum_{j=1}^q (r_{ij} - \bar{r}_i)^2 / (q-1), \text{ where } \bar{r}_i = \sum_{j=1}^q r_{ij} / q, \text{ and } r_{ij} = y_{ij} / \bar{y}_{.j}$$

This statistic can be see to measure Type-II stability.

Grouping of the indices and their similarity

The first nine statistics are based either on the deviation from average genotype effect (DG) = $y_{ij} - \bar{y}_i$ or on the G x E interaction term $l_{ij} = y_{ij} - \bar{y}_i - \bar{y}_{.j} + \bar{y}_{..}$ (in form of their sums of squares SS , regression coefficient or deviation from regression) and were classified into four groups (A, B, C, D):

Group A: DG, SS:	(S_i^2, CV_i) – Type I stability	
Group B: GE, SS:	$(\theta_i, \theta_{(i)}, W_i^2, \sigma_i^2)$ – Type	II
stability		
Group C: DG or GE regression coefficient :	(b_i, β_i) – Type II stability	
Group D: DG or GE regression deviation.	(δ_i^2) – Type III stability	

Lin et al (1986) noted:

(i) Since $\text{Var}(\log(y)) \sim \text{Var}(y) / (\text{mean}(y))^2 = (CV(y))^2$. Thus the two statistics in group A are equivalent, except for data transformation.

- (ii) The four statistics in group B are equivalent for the purpose of ranking genotypes. σ_i^2 , also is an unbiased estimate of variance of genotype i. An approximate test for homogeneity of σ_i^2 has been given by Shukla (1972b).
- (iii) Since $\beta_i = b_i - 1$, the two statistics in group C are equivalent. Similarly, the statistics of group D are equivalent.
- (iv) When variability in response can be satisfactorily expressed by a regression model, the regression coefficient (of group C) can serve as stability parameters and could be preferred to variability measures (group B) since, they (of Group C measures) provide information on shape of response along with its variation.

Stability indices and stability measures

- (i) The statistics in Group A measure Type I stability; those of Group B and Yau's index measure Type II while those of Group D measure Type III. The statistics of Group C are of Type I or Type II stability measures depending on the nature of the stable genotype. If stable genotype are defined by having $b_i=1$ ($\beta_i=0$) Type II is implied; but if they are defined by $b_i=0$ ($\beta_i=-1$), then Type I is implied.
- (ii) Type I stability: Type I indicates homeostasis, a biological concept (Becker 1981). It differs from agronomic concept of stability given by Type II. Although Type I is theoretically sound, but breeder do not use it frequently, for a breeder would like to select cultivars with high yields besides having Type I stability. Type I stability is associated with relatively poor yield in environments which are high yielding for other cultivars. Also, b_i and yield are positively correlated (Finlay and Wilkinson, 1963). Although wide (broad) adaptation may be desirable but difficult to achieve in practice. A still more convenient way would be to breed cultivars with (specific) adaptation to different environments to maximize the production. Since Type I stability does not depend on the presence of other genotypes, it has broad inference base. However, it does not provide information on the response structure.
- (iii) Type II stability: The inferences from Type II stability measures are relative to the genotypes included in the test. For example, a genotype A may be assessed stable and B unstable if A resembles majority of genotypes in the set more closely than does B. In another set of genotypes, if B resembles majority of genotypes more closely than does A, then B is stable and A unstable. This measure is useful for comparing a specific set of genotypes and thus, does not have a broad inference base for general assessment.
- (iv) Type III stability: Eberhart and Russell (1966) suggested another measure of stability based on mean square of deviation (σ_i^2). Thus, there are two measures of stability, (b_i , σ_i^2) for a single character. Use of σ_i^2 was advocated by Breese (1969) as he considered

'stability' should refer the unpredictable variation (irregularities) in response to environment. The variability of response to environment can be divided into predictable variation (given by regression mean squares) and unpredictable variation (measured by deviation MS, σ_i^2). This argument is sound but the measure of stability by deviation MS is inappropriate as it represent the goodness of fit of the model we choose. To support the argument in practice, one must measure independent variables explaining environment and a prediction model be made with them. The environmental index (based on means of all genotypes) can not provide an independent measure of environment potential. Thus low value of percent variance accounted for or high σ_i^2 or heterogeneous MS simply indicate that regression model is not adequate for stability and some other methods should be investigated. Type III is useful only when the prediction model is considered and is based on independently measured environmental variables.

3.2 Non-parametric Approach

In sequel to our previous discussion, we now consider the non-parametric approach to study genotype x environment interaction. This approach is used to search pattern in the genotypes and or in the environments.

The statistics under parametric approach express multivariate information (responses over multi-environments are considered multi-variate) in terms of a univariate, and measure only individual aspects (Types I, II, or III) of stability. It is possible to arrive at a contradiction, i.e. a genotype may be found stable for one type of stability measure but could be found unstable for the other measure(s). These do not provide any interrelationship among the genotype exploiting the response patterns from these (common) environments. Classifying genotypes into quantitatively homogeneous stability subsets, based on similarity of their responses to the environments, is another line of thought to evaluate interrelationships among the genotypes, and such an approach is considered to be non-parametric. The classification method has an advantage in the sense that although the genotypes are grouped on the basis of a specific data set, the relative relationship among genotypes can be independent of it or any specific data set. For instance, two genotypes say A and B, with dissimilar response patterns (unrelated) can always be grouped into two different stability sets, irrespective of the presence of genotypes resembling A and or B. We now consider methods for classification.

3.2.1 Cluster Analysis

Several commonly used methods for clustering genotypes (or environments) based on similarity of response characteristics are available in references cited in the end of this material. Every clustering technique has two considerations. (i) a definition of the similarity matrix, (ii) a strategy for grouping. Two cases arise. In one case, similarity is based on genetic effect and $G \times E$ interaction (means of $p \times q$ table) while in the other case similarity is based on $G \times E$ interaction only. Strategies for grouping could be, for instance incremental sums of squares (ISS) fusion strategy and group average (GA) fusion strategy (see, Cormak, 1971). This is the case of hierarchical clustering.

Limitations of clustering methods

The particular choices of similarity matrices and clustering strategies give rise to different cluster groups and this may lead to problems of preferring one method of cluster to other. Another criticism of clustering method is that it can also force unwanted structure on a data set suggesting misleading results.

2.2.2 Non-hierarchical clustering

In non-hierarchical clustering, the purpose is to group the units (genotypes or locations) in a number of disjoint classes chosen in advance using the information on a number of variables on them. The units with a class are expected to be homogeneous on the basis of some criterion. In the hierarchical clustering one can cut the dendrogram at a level of similarity to provide a selected number of groups but the statistical properties of such a grouping is not yet clear. In non-hierarchical clustering the groups of units are obtained by optimizing the selected criterion. Some of these are

- i. maximization of between-group sum of squares
- ii. maximal predictive classification,
- iii. minimizing the determinant of the pooled within-class dispersion matrix,
- iv. maximizing the total Mahalanobis squared distance between the groups.

2.2.3 Ordination techniques

Ordination techniques are used to simplify multivariate data for a set of individuals by summarizing relationships among individuals or among attributes describing them. This is done by producing a simple visual representation of the individuals as points which can be plotted to portray their relationships acceptably free of distortion. The ordination techniques try to reduce the dimensionality of the multivariate systems efficiently to

preserve the relationships among individuals as far as possible, but to provide a simplified view of those relationships in fewer dimensions than specified by original variables. There are two methods of ordination.

3.2.4 Principal Component Analysis (PCA).

PCA considers finding a new set of coordinate axes which accounts more effectively for the variation among individuals than do those based on original variables. PCA represent a transformation of data from one set of coordinate to another. This may not necessarily lead to reduction of dimensionality. However, when only (first) few principal components account for most of the variation, then it becomes effectively useful. Algebraically, the principal axes are determined by the latent vectors from the matrix of corrected sums of squares and products among variables. Elements of each vector specify the linear combinations of original variables necessary to give the corresponding PC and the associated latent root give the variation attributable to the component. PCA can also be applied on environment in same way as it could be done to genotypes. Mandel (1969) considered it for G x E interaction effect.

3.2.5 Principal Coordinate Analysis of genotypes. (PCO)

PCO analysis requires finding a set of rectangular coordinate axes which accounts as efficiently as possible for variation among individuals and may lead subsequently to a reduction in dimensionality for simplification. These objectives are similar to PCA but PCO is based on a much more general approach. It does not automatically assume that original variables define a multidimensional Euclidian space, in which relationship between pairs of individuals are indicated by Euclidian distance. Many similarity measures (e.g. correlation coefficients) or dissimilarity measures (distance) could be used. PCO involves two steps for computation.

- 1) presentation of the set of individuals as points in a coordinate space derived from the original matrix of measures. Gower (1966) showed that the interpoint Euclidian distances in this space are a simple function of the original measures of relationship between individuals. The significance of this method is that it refers individuals to Euclidian coordinate axes even when an initial coordinate framework is unavailable, and it represents original measures of relationships as Euclidian distances even if they are non-Euclidian.

2) carrying out a PCA on the data derived in step (1).

The two steps of PCO combined in one are given by Gower (1967). The only requirement to guarantee a distortion-free representation by PCO is that the original matrix of measures must be symmetric (so that no negative latent roots are obtained). In general, principal axes will not be a linear combination of original variables as in usual PCA. However, it is possible to investigate the relationship of original variables to each principal axis by correlating the set of principal coordinate scores for each axis with each of original variables. A correlation of large magnitude for a particular variable implies that it is strongly reflected in the axis concerned. Gower (1966) also showed that PCA is a special case of PCO when measures used in PCO are squared Euclidian distances.

4. Partitioning of GxE interaction

We may present

- a) the results of cluster analysis employed for zoning the environments and grouping the genotypes.
- b) partitioning of the GxE interaction using these groupings.

The care must be taken in justifying the groups resulting from a methods in terms of the number of groups, and the nature of locations and genotypes within. It is recommended that the groupings must be looked into the light of some other (independent) variables reflecting the physical properties of the environments and phenological and morphological traits of the genotypes.

A complete hierarchy should be presented with help of dendrogram when agglomerative methods of forming groups are used. Let n^e and n^g be the number of environment groups and number of genotypes groups respectively. Also let n_i^g be the number of genotypes in the i -th genotype group ($i=1...n^g$) and n_j^e be the number of environments in the j -th environment group ($j=1...n^e$). Note that $\sum_{j=1}^{n^e} n_j^e = q$ and $\sum_{i=1}^{n^g} n_i^g = p$. Further, with the reduced G x E data matrix one may present:

- (i) Partitioning of the variation related to grouping model.

Analysis of variance skeleton.

Source	d.f.	SS	MS
Environments (E)	$q-1$		
Among E groups	n^e-1		
Within E groups	$\sum(n_j^e-1)$		
Genotypes (G)	$p-1$		
Among G groups	n^g-1		
Within G groups	$\sum(n_i^g-1)$		
G x E	$(q-1)(p-1)$		
Among G groups x among E groups	-		
Among G groups x within E groups	-		
Within G groups x among E groups	-		
Within G groups x within E groups	-		
Residual	-		

(ii) Group performance plots and

(iii) Patterns of (G x E) interactions on grouped sets.

5. Stochastic Dominance of Varieties

This procedure emphasizes the riskiness of (new) genotype or variety. New crop varieties (or new technologies, in general) may often be regarded by farmers more risky than traditional ones. Risk may, therefore tend to act as an impediment to their adoption. Improved varieties that would be preferred by "risk-averse" farmers can be identified by stochastic dominance procedure under certain assumptions. Anderson (1974) used this procedure for analyzing data from the Sixth International Spring Wheat Yield Nurseries administered by CIMMYT. He made following three assumptions.

- (i) it makes sense to talk about (or large regional) probability distribution of wheat yields,
- (ii) the selection of sites, cooperators, fields and growing and disease conditions is representative of the relevant world (or regional) domain of production, and
- (iii) yield per se provides a reasonable surrogate for the argument of the average farmer's utility function.

Menz (1980) used cluster analysis of Byth et al (1976) to analyse CIMMYT International Spring Wheat Yield Nurseries over five years and also used stochastic dominance. He found considerable degree of agreement in the results based on the two methods.

6. Additive Main Effects and Multiplicative Interaction Model

The AMMI model stands for additive main effects and multiplicative interaction model. The data on GxE are fitted using

- i. main effects of genotypes and environments,
- ii. the interaction GxE is fitted as sum of multiplicative PCA scores for genotype and environments.

Main advantage of this method is that it facilitates examination of the pattern of GxE interaction as expressed by a general number of principal components. Further details are available in a recent series of articles including Gauch (1988), and Gauch and (1988).

7. Inter-site Transferability of Crop Varieties

Development of varieties and their evaluation often takes place on a limited range of environments (e.g. experimental stations) but they are actually targeted for production in much larger set of environments (e.g. farmers' field). Therefore, transferability of variety response to a new location is an important aspect of variety recommendation. Singh et al (1996) provided a statistical measure of the transferability of a variety using multi-

locational data. The approach is as follows. For a given variety say i , its response to the environment can be modeled as a linear regression on environmental index (often considered to be sound biological measure and is taken as mean of all the genotypes at that location). To evaluate transferability of the genotype response to an environment say j -th, fit the linear regression of yield on environmental index using data on (response, index) pairs for all locations except the j -th location and compute the difference in yield response observed and predicted response at the j -th location using the above linear regression. Such a difference has been called inter-site residual (Wood and Cady, 1981) and predicted residual (Cook and Wiesberg 1982). Such differences can be obtained by leaving one location at a time. Their (weighted) sum of squares gives inter-site residual sum of squares. For assessing the inter-site residuals, we may consider plot- residuals as within-site residuals. A measure (P) of transferability for the genotype (i under consideration) then is the ratio of inter-site transfer residual sum of square to within-site residual sum of squares weighted with replications. Statistical distributions of linear functions of P has been worked out when error variances over locations are homogeneous/ heterogeneous. Six trials with number of locations varying from 16-53 and variety varying from 21-23 have been presented for barley and wheats in Singh et al (1996).

8. An Illustration

We list in the following printout from a GENSTAT 5 program written for analyzing data from multi-locational variety trials conducted in randomized complete block designs. The program codes are available for designs in complete blocks as well as in incomplete blocks on diskette.

```

3  "
-4
-5
-6  *****
-7      GENSTAT program for analyzing multi-locational variety trials conducted
-8      in complete blocks. Data from all locations are in a single file.
-9
-10     This includes
-11     i.    analysis of data from individual environments
-12     ii.   tests for homogeneity of error variances,
-13     iii.  combined analysis of data for GxE interaction
-14          under homogeneous/heterogeneous errors
-15     iv.   tests for parallelism of regression lines
-16     v.    common stability statistics
-17     vi.   hierarchical cluster analysis of genotypes
-18     vii.  hierarchical cluster analysis of environments
-19     viii. clustering of genotypes and environments into groups which
maximizes
-20          GxE interaction between the groups of genotype and
-21          groups of environments (Corsten and Denis 1990)

```

```

-22          ix.    non-hierarchical cluster analysis of genotypes
-23          x.    non-hierarchical cluster analysis of environments
-24          xi.   principal component analysis of genotypes
-25          xii.  principal component analysis of environments
-26          xiii. heritability of the traits
-27
-28          Software: GENSTAT 5 Rel 4.1
-29
-30          .....
-31          "
-32
-33
-34
-35
-36          Open ch=2; fi=in ; "Give the name of the text data file"   Name = 'mlvt1_2.txt'
-37
-38
-39
-40
-41          Scal Alpha; 0.05
-42          Scal NRoots;3      " for PCA "
-43          Scal GxEI%; 60    " % of GxEI explained by between G-group and E-group"
-44          Scal NLoc, NRepMax, NGeno
-45          Skip[ch=2]2 : Read [ch=2] NLoc,NRepMax,NGeno

          Identifier   Minimum      Mean      Maximum      Values      Missing
          NLoc          10.00        10.00      10.00         1           0
          NRepMax       3.000        3.000      3.000         1           0
          NGeno         15.00        15.00      15.00         1           0

-46          Fact[leve=NLoc] Loc : Fact[leve=NRepMax]Rep : Fact[leve=NGeno]Geno
-47          Skip[ch=2]1 : Read[ch=2] Loc, Rep, Geno, Yield : Clos 2

          Identifier   Minimum      Mean      Maximum      Values      Missing
          Yield        150.0        3692      9000         450          1

          Identifier   Values      Missing      Levels
          Loc          450          0           10
          Rep          450          0           3
          Geno         450          0           15

-48
-49          Scal NRep[1...NLoc] " replications under individual environments "
-50
-51          For i=1...NLoc ; dr=NRep[1...NLoc]
-52          Rest Rep ; Loc==i : Calc dr=Max(Rep): Rest Rep : Endf
-53
-54
-55          Scal NRepAvrg : Calc NRepAvrg=VMean(!p(NRep[1...NLoc]))
-56          Scal NObs, NGxNL : Calc NObs=NGeno*VSum(!p(NRep[1...NLoc]))
-57          Calc NGxNL=NGeno*NLoc : Prin NGxNL,NObs ; deci=0

          NGxNL      NObs
          150        450

-58
-59          Units[NObs]
-60
-61          " ===== Below is only for statistical programmers use ===== "
-62          " ===== Below is only for statistical programmers use ===== "
-63          " ===== Below is only for statistical programmers use ===== "
-64
-65
-66          " 1. Individual locations analysis "
-67          " 1. Individual locations analysis "
-68
-69          Scal sigma2, ss,df
-70          Vari[Nval=NGeno] Mean[1...NLoc], GenoMean
-71          Vari[Nval=NLoc]ErrMS,ErrDF, Weight
-72
-73          Vari[Nval=NGeno] Mean[1...NLoc], GenoMean
-74          Vari[Nval=NLoc]CV%,SEM, LocMean
-75          Fact[Leve=NLoc; Valu=1...NLoc] LocNum
-76
-77          Block Rep/Geno : Treat Geno

```

```

78 For I=1...NLoc      ; MN=Mean[1...NLoc]
79 Print '***** Location number is = ', I, '*****'
80 Rest Yield; Cond=Loc.EQ.I
81 Anova[prin=a; fpro=y] Yield
82 Akeep Rep.Geno; ss=ss; df=df
83 Akeep Geno; Means=TDum
84 Calc sigma2=ss/df
85 Calc ErrMS[I]=sigma2 : Calc ErrDF[I]=df : Calc Weight[I]=NRep[I]/sigma2
86 Equa TDum;MN
87 Dele[Rede=Y] TDum
88 Calc LocMean[I]=Mean(Yield)
89 Calc CV[I]=100*Sqrt(sigma2)/Mean(Yield)
90 Calc SEM[I]=sqrt(sigma2/NRep[I])
91
92 Rest Yield
93
94 Endf

```

```

***** Location number is =          1.000 *****

```

```

94.....

```

```

***** Analysis of variance *****

```

```

Variate: Yield

```

Source of variation	d.f.	s.s.	m.s.	v.r.	F pr.
Rep stratum	2	27563.	13781.	0.23	
Rep.Geno stratum					
Geno	14	666583.	47613.	0.78	0.682
Residual	28	1711141.	61112.		
Total	44	2405287.			

```

***** Location number is =          2.000 *****

```

```

94.....

```

```

***** Analysis of variance *****

```

```

Variate: Yield

```

Source of variation	d.f.	s.s.	m.s.	v.r.	F pr.
Rep stratum	2	165032.	82516.	0.18	
Rep.Geno stratum					
Geno	14	5153189.	368085.	0.79	0.672
Residual	28	13060300.	466439.		
Total	44	18378521.			

```

.... other locations' ANOVA dropped .....

```

```

***** Location number is =          10.00 *****

```

```

94.....

```

```

***** Analysis of variance *****

```

```

Variate: Yield

```

Source of variation	d.f.	s.s.	m.s.	v.r.	F pr.
Rep stratum	2	933760.	466880.	0.47	
Rep.Geno stratum					
Geno	14	9569067.	683505.	0.69	0.769
Residual	28	27909973.	996785.		
Total	44	38412800.			

95

96 Prin LocNum,LocMean,CV%,SEM,ErrMS,ErrDF; fiel=9

LocNum	LocMean	CV%	SEM	ErrMS	ErrDF
1	1409	17.54	142.7	61112	28.00
2	6069	11.25	394.3	466439	28.00
3	5417	14.43	451.4	611252	28.00
4	5324	24.61	756.6	1717324	28.00
5	4086	20.23	477.3	683349	28.00
6	1264	47.34	345.4	357816	28.00
7	2816	13.59	221.0	146458	27.00
8	415	20.48	49.1	7220	28.00
9	5872	18.84	638.7	1223871	28.00
10	4227	23.62	576.4	996785	28.00

97 Hist[ngroup=5] CV% :& ErrMS : & LocMean

Histogram of CV%

-	16	3	***
16 -	24	5	*****
24 -	32	1	*
32 -	40	0	
40 -		1	*

Scale: 1 asterisk represents 1 unit.

Histogram of ErrMS

-	400000	4	****
400000 -	800000	3	***
800000 -	1200000	1	*
1200000 -	1600000	1	*
1600000 -		1	*

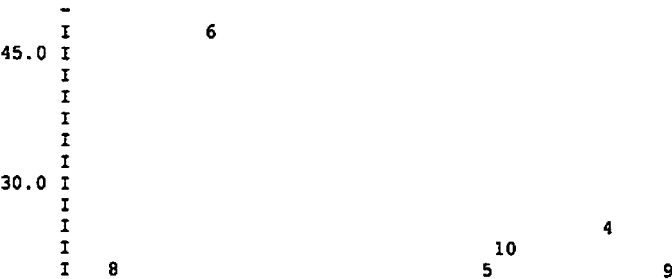
Scale: 1 asterisk represents 1 unit.

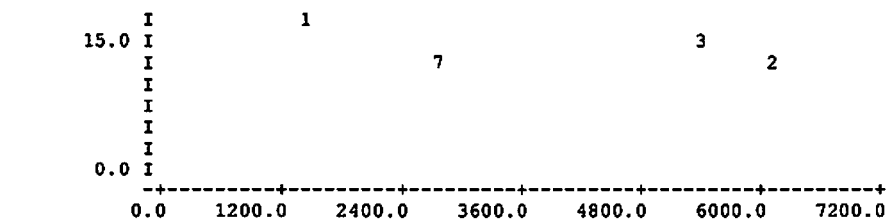
Histogram of LocMean

-	1500	3	***
1500 -	3000	1	*
3000 -	4500	2	**
4500 -	6000	3	***
6000 -		1	*

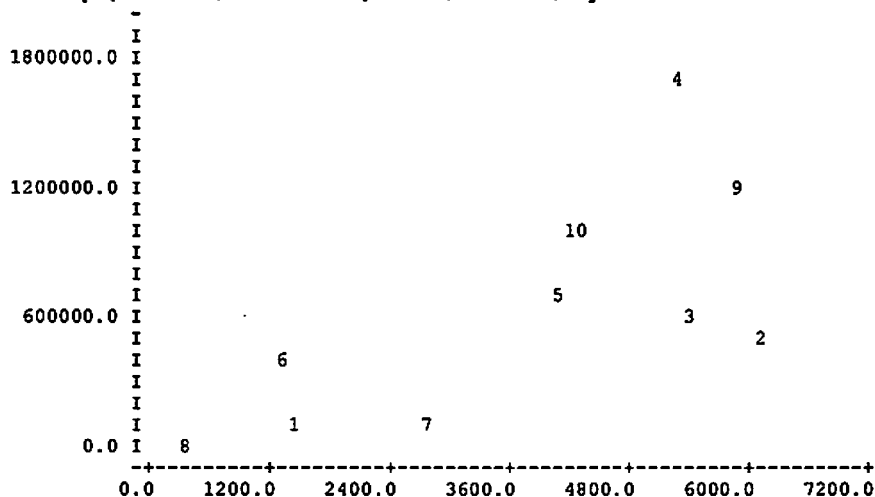
Scale: 1 asterisk represents 1 unit.

98 Graph[nrows=20; ncolumn=60] CV%; LocMean; symb=LocNum





CV% v. LocMean using factor LocNum
99 Graph[nrows=20; ncolumn=60] ErrMS; LocMean; symb=LocNum



ErrMS v. LocMean using factor LocNum

```

100
101 " 2. Bartlette Test for homogeneity of error variances "
102 " 2. Bartlette Test for homogeneity of error variances "
103 Scal Prob, PoolMS, PoolDF
104 Calc PoolDF=Sum(ErrDF): & PoolMS=Sum(ErrMS*ErrDF)/PoolDF
105 Calc Prob=( PoolDF*Log(PoolMS)-Sum(ErrDF*Log(ErrMS)) )/ \
106 ( 1+ (Sum(1/ErrDF)-1/PoolDF)/3/(NLoc-1) )
107 Calc Prob=Cuchi(Prob;NLoc-1)
108 Prin PoolMS,PoolDF,Prob

      PoolMS      PoolDF      Prob
      626886      279.0      0

109 If Prob.lt.Alpha
110 Print ' Location error variances heterogeneous at ' , Alpha, ' probability'

      Alpha
      0.05000 probability
111 Else
112 Print ' Location error variances homogeneous at ' , Alpha, ' probability'
113 Endif
114
115 Vari[Nval=NLoc] LocMean
116
117 " 3. Combined analysis over locations "
118 " 3. Combined analysis over locations "
119
120 Bloc Loc.Rep/Geno : Trea Loc*Geno
121 Anova[prin=a,m; pse=m; pfact=1; fpro=y] Yield

```

121.....

***** Analysis of variance *****

Variate: Yield

Source of variation	d.f. (m.v.)	s.s.	m.s.	v.r.	F pr.
Loc.Rep stratum					
Loc	9	1.759E+09	1.955E+08	148.92	<.001
Residual	20	2.625E+07	1.313E+06	2.09	
Loc.Rep.Geno stratum					
Geno	14	2.097E+07	1.498E+06	2.38	0.004
Loc.Geno	126	1.193E+08	9.470E+05	1.51	0.003
Residual	279(1)	1.755E+08	6.289E+05		
Total	448(1)	2.101E+09			

***** Tables of means *****

Variate: Yield

Grand mean 3691.

Loc	1	2	3	4	5	6	7
	1409.	6069.	5417.	5324.	4086.	1264.	2821.
Loc	8	9	10				
	415.	5872.	4227.				
Geno	1	2	3	4	5	6	7
	3167.	3751.	3744.	3737.	4019.	3599.	3784.
Geno	8	9	10	11	12	13	14
	3584.	3713.	3635.	3694.	4122.	3810.	3542.
Geno	15						
	3458.						

*** Standard errors of means ***

Table	Loc	Geno
rep.	45	30
d.f.	20	279
e.s.e.	170.8	144.8

(Not adjusted for missing values)

```

122
123 Fact[Leve=NGeno; Valu=1...NGeno] GenoNum
124 Fact[Leve=NLoc; Valu=1...NLoc] LocNum
125 Calc GenoMean=VMean(!P(Mean[1...NLoc]))
126 For i=1...NLoc : Calc LocMean$(i)=Mean(Mean[i]) : Endf
127
128 Prin GenoNum,Mean[1...NLoc],GenoMean; field=7

```

GenoNum	Mean[1]	Mean[2]	Mean[3]	Mean[4]	Mean[5]	Mean[6]	Mean[7]	Mean[8]	Mean[9]
Mean[10]									
1	1167	6178	4000	5167	4026	660	2107	441.7	3567
2	1333	6044	4933	6635	3173	1457	2308	391.7	6300
3	1341	6556	6100	5967	3423	1050	2611	316.7	5813
4	1411	5700	5292	6245	4314	699	2785	283.3	6958
5	1464	5511	5867	5617	5551	1437	2833	333.3	7558
6	1511	5595	6050	5489	3699	1168	3063	275.0	6003
7	1411	5933	5350	5737	4987	1451	2740	358.3	5428
8	1495	5867	4700	5445	3070	987	3181	375.0	6837
9	1516	6389	5133	5700	3897	1899	2596	491.7	5387
10	1576	5611	5167	5056	3423	1565	3061	408.3	6200
11	1511	6189	5400	3833	5135	1254	3048	575.0	5637
12	1161	6455	6600	5958	5019	1391	3501	483.3	6280
13	1357	6389	6117	4278	4750	1469	2845	641.7	6360
14	1533	6456	4800	4481	3494	1264	2816	366.7	5200
15	1354	6167	5750	4259	3333	1203	2817	483.3	4558

GenoMean
3167
3751

```

3744
3737
4019
3599
3784
3584
3713
3635
3694
4122
3810
3542
3458

129 Prin[orie=a] LocMean ;fiel=7

      LocMean  1409   6069   5417   5324   4086   1264   2821   415   5872

      LocMean  4227

130
131 " Transpose data matrix for the sake of convenience"
132 Matr [Rows=NGeno; Colu=NLoc] GE : & [Rows=NLoc; Colu=NGeno] EG
133 Equa !P(Mean[1...NLoc]) ; EG
134 Calc GE=Tran(EG)
135
136 Vari[Nval=NLoc] GMean[1...NGeno]
137 Equa GE; !P(GMean[1...NGeno])
138
139      Dele GE, EG
140
141 Vari[ Nval=NGxNL] GEData
142 Equa !P(Mean[1...NLoc]); GEData
143
144 " ***** Analysis of variance *****"
145
146 Fact[ Leve=NGeno; Nval=NGxNL] Genol : Fact[ Leve=NLoc; Nval=NGxNL] Loc1
147 Gene Loc1, Genol
148 Bloc
149 Treat Loc1*Genol
150 Anov[prin=a;fpro=y] GEData

150.....

**** Analysis of variance ****

Variate: GEData

Source of variation      d.f.      s.s.      m.s.      v.r.  F pr.
Loc1                      9 5.8640E+08 6.5156E+07
Genol                     14 6.9898E+06 4.9927E+05
Loc1.Genol                126 3.9774E+07 3.1566E+05
Total                     149 6.3317E+08

151
152
153 Vari[Nval=NGxNL]AllWet: Equa !p(#NGeno(#Weight)); AllWet
154 Anov[weight=AllWet; prin=a;fpro=y] GEData

154.....

**** Analysis of variance ****

Variate: GEData
Weight variate: AllWet

Source of variation      d.f.      s.s.      m.s.      v.r.  F pr.
Loc1                      9 9442.689 1049.188
Genol                     14 61.356 4.383
Loc1.Genol                126 201.326 1.598
Total                     149 9705.370

```

```

155 Dele[Rede=Y]Loc1,Genol : dele[rede=y] AllWet
156
157 " Note--- Mean[1...NLoc] of length NGeno and GMean[1...NGeno] of length NLoc "
158
159 " 4. Partition GxE Int in heterogeneity of linear regressions"
160 " 4. Partition GxE Int in heterogeneity of linear regressions"
161
162 " 4.1 Test for heterogeneity of linear regressions: unweighted analysis "
163 Fact[ Leve=NGeno; Nval=NGxNL] Genol : Fact[ Leve=NLoc; Nval=NGxNL] Loc1
164 Gene Loc1, Genol
165
166 Bloc Loc.Rep/Geno : Trea Geno*Pol(Loc;1;LocMean) : Anov(prin=a;fpro=y)Yield
166.....

**** Analysis of variance ****

Variate: Yield

Source of variation      d.f. (m.v.)      s.s.      m.s.      v.r.      F pr.

Loc.Rep stratum
Loc                      9      1.759E+09  1.955E+08  148.92  <.001
  Lin                   1      1.759E+09  1.759E+09  1340.28  <.001
  Deviations           8      0.535E-21  0.669E-22   0.00   1.000
Residual                20      2.625E+07  1.313E+06   2.09

Loc.Rep.Geno stratum
Geno                    14      2.097E+07  1.498E+06   2.38   0.004
Loc.Geno                126     1.193E+08  9.470E+05   1.51   0.003
  Lin.Geno              14      1.147E+07  8.192E+05   1.30   0.205
  Deviations           112     1.079E+08  9.630E+05   1.53   0.003
Residual                279(1)    1.755E+08  6.289E+05

Total                   448(1)    2.101E+09

167
168
169 Vari[Nval=NGxNL]AllLoc : Equa !P(#NGeno(#LocMean)); AllLoc
170
171 Model GEData : Fit[Prin=m,s,a;fpro=yes]AllLoc+Genol+AllLoc.Genol

171.....

**** Regression Analysis ****

Response variate: GEData
  Fitted terms: Constant + AllLoc + Genol + AllLoc.Genol

*** Summary of analysis ***

      d.f.      s.s.      m.s.      v.r.      F pr.
Regression      29    5.972E+08    20593622.    68.74  <.001
Residual       120    3.595E+07     299591.
Total          149    6.332E+08    4249436.

Percentage variance accounted for 92.9
Standard error of observations is estimated to be 547.
* MESSAGE: The following units have large standardized residuals:
      Unit      Response      Residual
      56         3833.         -2.73
      121        3567.         -3.13
* MESSAGE: The error variance does not appear to be constant:
      large responses are more variable than small responses

*** Accumulated analysis of variance ***

Change      d.f.      s.s.      m.s.      v.r.      F pr.
+ AllLoc      1    5.864E+08    5.864E+08    1957.35  <.001
+ Genol       14    6.990E+06    4.993E+05     1.67   0.072
+ AllLoc.Genol 14    3.823E+06    2.731E+05     0.91   0.549

```

Residual	120	3.595E+07	2.996E+05
Total	149	6.332E+08	4.249E+06

```

172 "Fit[Prin=*,Cons=o,fpro=yes; tpro=y]Genol/AllLoc"
173
174 "4.2 Test for heterogeneity of linear regressions: weighted analysis "
175 Vari[Nval=NGxNL]AllWet: Equa !p(#NGeno(#Weight)); AllWet
176 Model[weight=AllWet;disp=1]GEData
177 Fit[Prin=m,s,a,fpro=yes]AllLoc+Genol+AllLoc.Genol

```

177.....

***** Regression Analysis *****

Response variate: GEData
Weight variate: AllWet
Fitted terms: Constant + AllLoc + Genol + AllLoc.Genol

*** Summary of analysis ***

	d.f.	s.s.	m.s.	v.r.	chi pr
Regression	29	9539.8	328.959	328.96	<.001
Residual	120	165.6	1.380		
Total	149	9705.4	65.137		

* MESSAGE: ratios are based on dispersion parameter with value 1

Percentage variance accounted for 97.9

Standard error of observations is fixed at 1.00

* MESSAGE: The following units have large standardized residuals:

Unit	Response	Residual
12	1161.00	-3.03
16	6178.00	3.03
20	5511.00	-2.82
65	5551.00	2.72

* MESSAGE: The following units have high leverage:

Unit	Response	Leverage
106	441.67	0.92
107	391.67	0.92
108	316.67	0.92
109	283.33	0.92
110	333.33	0.92
111	275.00	0.92
112	358.33	0.92
113	375.00	0.92
114	491.67	0.92
115	408.33	0.92
116	575.00	0.92
117	483.33	0.92
118	641.67	0.92
119	366.67	0.92
120	483.33	0.92

*** Accumulated analysis of variance ***

Change	d.f.	s.s.	m.s.	v.r.	chi pr
+ AllLoc	1	9442.689	9442.689	9442.69	<.001
+ Genol	14	61.356	4.383	4.38	<.001
+ AllLoc.Genol	14	35.762	2.554	2.55	0.001
Residual	120	165.563	1.380		

Total	149	9705.370	65.137
-------	-----	----------	--------

* MESSAGE: ratios are based on dispersion parameter with value 1

```

178 "Fit[Prin=*,Cons=o,fpro=yes; tpro=y]Genol/AllLoc"
179
180 dele AllLoc, AllWet
181
182 " 5. Compute stability indices "
183 " 5. Compute stability indices "

```

```

184
185 Vari[nvalu=NGeno] GenoCV
186 Calc GenoCV=100.*Sqrt(Vvar(!P(Mean[1...NLoc])))/GenoMean
187
188 Vari[ Nval=NGeno] Slope,SeSlop,DeviMS, Wricke,Pla_Pet,Plaisted, Shukla, YauH
189 Vari[Nval=NGeno] DevRegDF, Probb1,ProbDev
190 Vari[ Nval=NGeno] SlopeW,SeSlopW, Probb1W,DeviSSW,ProbDevW, DevRgDFW
191 Vari[Nval=NGeno]RSq%,RSqW% " Goodness-of-fit % R-squares adjusted for df"
192 For I=1...NGeno ; Y=GMean[1...NGeno]
193 Scal YMeanSq, YMeanSqW : Calc YMeanSq=Var(Y)
194 Calc YMeanSqW=Sum(Weight*(Y-Sum(Weight*Y)/Sum(Weight))**2)/(Nval(Y)-1)
195
196 " Yau and Hamblin (1995)'s stability index"
197 Calc YauH$I=Var(Y/LocMean)
198
199 " Unweighted regression analysis"
200 Model Y ; Fitt=F
201 Fit[prin=*] LocMean
202 RKeep ; Est=Est; Se=Se ; Devi=SS ; DF=df
203 Calc Slope$I =Est ${2} : & SeSlop$I=Se${2} : & DeviMS $I =SS/df
204 Calc DevRegDF$I=df : & RSq$I=100*(1-SS/df/YMeanSq)
205 "
-206 Graph[nrows=20; ncolumn=60] Y,F; LocMean; symb='o','.'; Meth=p,c
-207 "
208
209 " Weighted regression analysis"
210 Model[Weight=Weight] Y ; Fitt=F
211 Fit[prin=*] LocMean
212 RKeep ; Est=Est; Se=Se ; Devi=SS ; DF=df
213 Calc SlopeW$I =Est ${2} : & SeSlopW$I=Se${2} : & DeviSSW$I =SS
214 Calc DevRgDFW$I=df : & RSqW$I=100*(1-SS/df/YMeanSqW)
215
216 Endf
217
218 Calc Probb1=Abs(Slope-1)/SeSlop : & Probb1=Cut(Probb1;DevRegDF)
219 Calc ProbDev=DeviMS/(PoolMS/NRepAvg)
220 Calc ProbDev=CuF(ProbDev;DevRegDF;PoolDF)
221 " above is based on an average number of replications. Use the weighted
-222 analysis results
-223 "
224
225 Calc Probb1W=Abs(SlopeW-1)/SeSlopW : & Probb1W=Cut(Probb1W;DevRgDFW)
226 Calc ProbDevW=Cuchi(DeviSSW;DevRgDFW)
227
228 " Get GxE interactions for stability indices"
229
230 Vari[ Nval=NGxNL] GEInt
231 Bloc Loc1.Genol
232 Trea Loc1+Genol
233 Anov[prin=*] GEData; Res=GEInt
234 AKee Loc1.Genol ; ss =GxEISS
235
236 Prin GxEISS

      GxEISS
39773742

237
238 Calc GEInt=GEInt*GEInt
239 Tabu [Class=Genol] GEInt; Tota=TDum
240
241 Equa TDum; Wricke
242 Dele[Rede=Y] TDum
243
244 Scal SsGE
245 Calc SsGE=Sum(Wricke)
246
247 Calc Pla_Pet=(NGeno*Wricke+SsGE)/(2*(NGeno-1)*(NLoc-1))
248 Calc Plaisted=(-NGeno*Wricke/(NGeno-1)+SsGE)/((NGeno-2)*(NLoc-1))
249 Calc Shukla = (NGeno*Wricke - SsGE/(NGeno-1))/(NGeno-2)/(NLoc-1)
250
251 " Correlation between indices "
252 Corr[Print=c] GenoMean, Slope, DeviMS,GenoCV,Wricke,Pla_Pet,Plaisted,Shukla, \
253 YauH, SlopeW,DeviSSW
*** Correlation matrix ***

```

GenoMean	1.000						
Slope	0.763	1.000					
DeviMS	-0.367	-0.175	1.000				
GenoCV	-0.142	0.508	0.442	1.000			
Wricke	-0.352	-0.133	0.987	0.484	1.000		
Pla_Pet	-0.352	-0.133	0.987	0.484	1.000	1.000	
Plaisted	0.352	0.133	-0.987	-0.484	-1.000	-1.000	1.000
Shukla	-0.352	-0.133	0.987	0.484	1.000	1.000	-1.000
YauH	-0.117	0.004	0.501	0.285	0.497	0.497	-0.497
SlopeW	0.876	0.794	-0.472	0.005	-0.443	-0.443	0.443
DeviSSW	-0.151	0.048	0.788	0.478	0.810	0.810	-0.810
GenoMean		Slope	DeviMS	GenoCV	Wricke	Pla_Pet	Plaisted
Shukla	1.000						
YauH	0.497	1.000					
SlopeW	-0.443	-0.319	1.000				
DeviSSW	0.810	0.306	-0.223	1.000			
Shukla		YauH	SlopeW	DeviSSW			

254 Print GenoNum, GenoMean, Slope,SeSlop,Probb1, DeviMS,ProbDev,RSq%, GenoCV,YauH;\n

255 field=8

GenoNum	GenoMean	Slope	SeSlop	Probb1	DeviMS	ProbDev	RSq%	GenoCV	YauH
1	3167	0.875	0.12606	0.1757	621262	0.0034	83.99	62.19	0.03663
2	3751	1.062	0.10712	0.2893	448611	0.0323	91.53	61.36	0.02380
3	3744	1.107	0.06156	0.0597	148129	0.6857	97.29	62.39	0.01585
4	3737	1.113	0.08666	0.1149	293599	0.1959	94.79	63.55	0.03950
5	4019	1.097	0.11631	0.2150	528832	0.0116	90.71	59.38	0.02798
6	3599	1.001	0.08054	0.4966	253568	0.2931	94.46	59.44	0.02250
7	3784	0.988	0.06195	0.4260	150046	0.6776	96.57	55.27	0.01113
8	3584	1.007	0.09379	0.4698	343905	0.1132	92.70	60.58	0.01921
9	3713	0.945	0.05525	0.1744	119352	0.8028	97.01	53.77	0.03144
10	3635	0.929	0.05189	0.1041	105270	0.8543	97.26	53.93	0.01285
11	3694	0.920	0.10187	0.2266	405679	0.0549	89.94	54.37	0.03123
12	4122	1.122	0.05947	0.0370	138238	0.7270	97.53	57.37	0.01526
13	3810	1.007	0.08812	0.4703	303553	0.1764	93.50	56.74	0.03940
14	3542	0.925	0.08468	0.2013	280336	0.2248	92.93	56.23	0.01360
15	3458	0.902	0.09509	0.1671	353476	0.1015	90.82	56.75	0.01741

256 Print GenoNum, GenoMean, SlopeW,SeSlopW,Probb1W, DeviSSW,ProbDevW, RSqW%; fiel=8

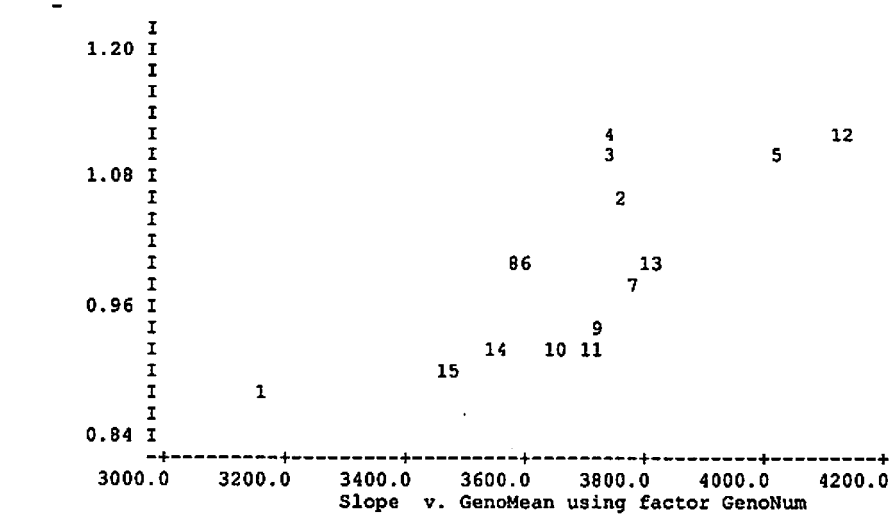
GenoNum	GenoMean	SlopeW	SeSlopW	Probb1W	DeviSSW	ProbDevW	RSqW%
1	3167	0.841	0.06573	0.0209	21.76	0.0054	94.76
2	3751	0.968	0.05322	0.2809	14.26	0.0751	97.34
3	3744	1.052	0.03527	0.0890	6.27	0.6175	99.00
4	3737	1.036	0.04158	0.2080	8.71	0.3677	98.57
5	4019	1.079	0.05987	0.1105	18.05	0.0208	97.30
6	3599	1.034	0.04771	0.2508	11.46	0.1768	98.11
7	3784	1.022	0.03232	0.2558	5.26	0.7294	99.11
8	3584	0.992	0.05240	0.4402	13.83	0.0863	97.54
9	3713	0.969	0.03524	0.2012	6.26	0.6187	98.82
10	3635	0.982	0.03735	0.3218	7.03	0.5338	98.71
11	3694	0.984	0.04245	0.3550	9.08	0.3359	98.35
12	4122	1.117	0.05394	0.0308	14.65	0.0663	97.94
13	3810	0.986	0.04448	0.3816	9.97	0.2674	98.20
14	3542	0.993	0.04376	0.4367	9.65	0.2908	98.28
15	3458	0.946	0.04308	0.1227	9.34	0.3141	98.16

257 Prin GenoNum, GenoMean,Wricke,Pla_Pet,Plaisted,Shukla; fiel=8

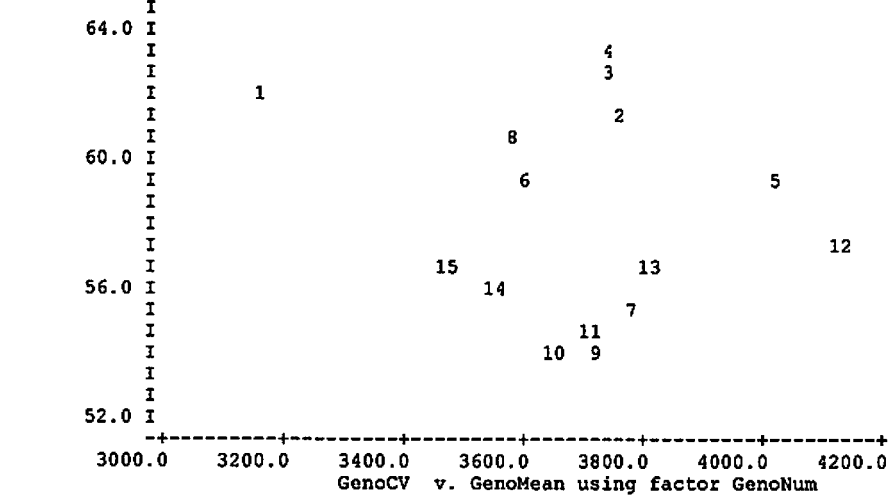
GenoNum	GenoMean	Wricke	Pla_Pet	Plaisted	Shukla
1	3167	5578639	489894	288860	690928
2	3751	3739272	380408	305704	455112
3	3744	1635091	255159	324973	185345
4	3737	2844787	327165	313895	340434
5	4019	4596131	431412	297857	564966
6	3599	2028564	278580	321370	235790
7	3784	1205943	229615	328903	130326
8	3584	2753349	321722	314733	328712
9	3713	1073016	221702	330120	113284
10	3635	1039503	219707	330427	108988
11	3694	3497666	366027	307917	424137
12	4122	1688955	258365	324480	192251
13	3810	2430222	302488	317692	287285
14	3542	2461612	304357	317404	291309

258 15 3458 3200992 348368 310633 386102

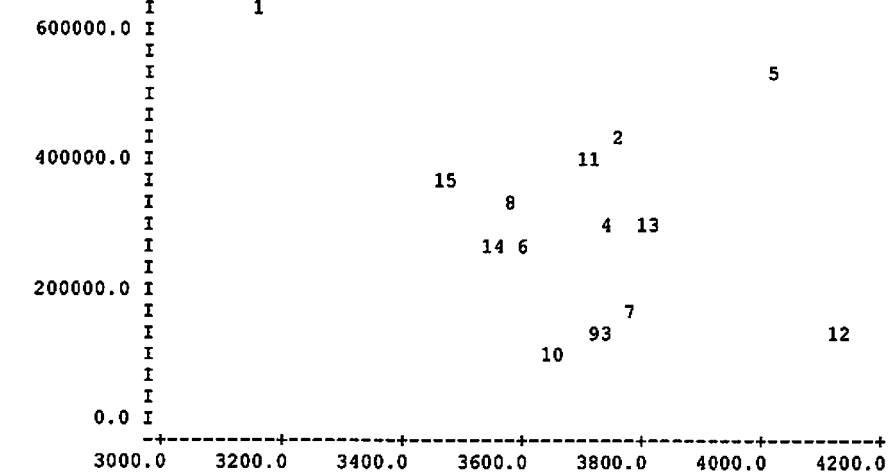
259 Graph[nrows=20; ncolumn=60] Slope; GenoMean; Symb=GenoNum



260 Graph[nrows=20; ncolumn=60] GenoCV; GenoMean; Symb=GenoNum

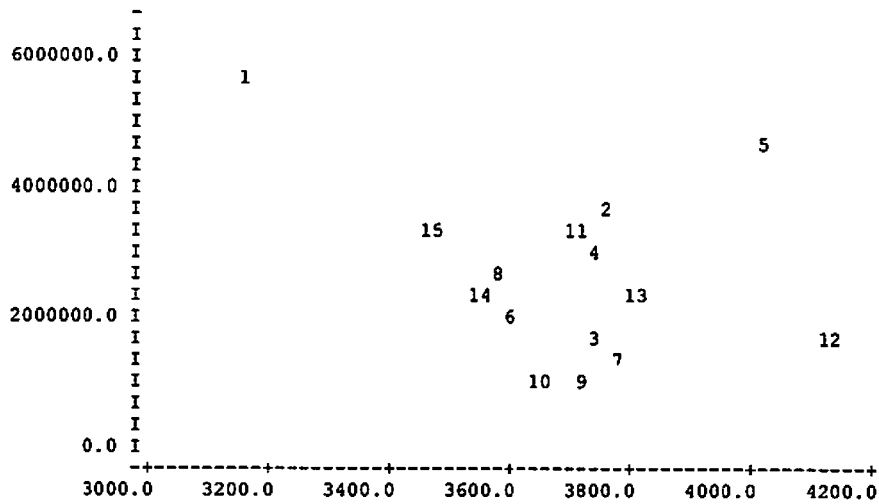


261 Graph[nrows=20; ncolumn=60] DeviMS; GenoMean; Symb=GenoNum



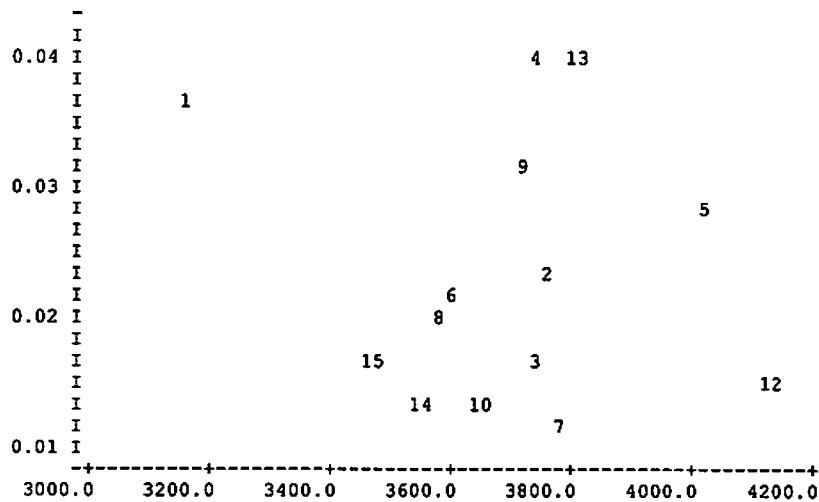
DeviMS v. GenoMean using factor GenoNum

262 Graph[nrows=20; ncolumn=60] Wricke; GenoMean; Symb=GenoNum



Wricke v. GenoMean using factor GenoNum

263 Graph[nrows=20; ncolumn=60] YauH; GenoMean; Symb=GenoNum



YauH v. GenoMean using factor GenoNum

Points coinciding with 9

11

```

264
265
266 Vari[nval=NGeno]RGenoMn,RSlope,RDeviMS,RGenoCV,RWricke, \
267   RPla_Pet,RPlaist, RShukla, RYauH
268 For D= GenoMean, Slope, DeviMS,GenoCV,Wricke,Pla_Pet,Plaisted, \
269   Shukla,YauH ; \
270   DD= RGenoMn, RSlope, RDeviMS,RGenoCV,RWricke, \
271   RPla_Pet,RPlaist,RShukla,RYauH
272 Vari[valu=1...NGeno]Order, DD
273 Sort[dire=d] D, Order

```

```

274 Sort Order, D, DD
275 endf
276
277
278
279 "          Correlations between ranks "
280 Corr[Prin=c] \
281   RGenoMn, RSlope, RDeviMS, RGenoCV, RWricke, RPla_Pet, RPlaist, RShukla, RYauH

*** Correlation matrix ***
  RGenoMn    1.000
  RSlope    0.739    1.000
  RDeviMS   -0.189   -0.207    1.000
  RGenoCV   -0.014    0.486    0.404    1.000
  RWricke   -0.200   -0.136    0.968    0.475    1.000
  RPla_Pet  -0.200   -0.136    0.968    0.475    1.000    1.000
  RPlaist    0.200    0.136   -0.968   -0.475   -1.000   -1.000    1.000
  RShukla   -0.200   -0.136    0.968    0.475    1.000    1.000   -1.000
  RYauH     0.018    0.046    0.493    0.304    0.507    0.507   -0.507

          RGenoMn  RSlope  RDeviMS  RGenoCV  RWricke  RPla_Pet  RPlaist
RShukla    1.000
RYauH      0.507    1.000

          RShukla  RYauH

282
283 dele RSlope, RDeviMS, RWricke, RPla_Pet, RPlaist, RShukla, RYauH
284 dele DevRegDF, Probl, ProbDev
285 dele SlopeW, SeSlopW, ProblW, DeviSSW, ProbDevW, DevRgDFW
286 dele RSq%, RSqW%
287
288 " 6. Hierarchical Clustering of Genotypes "
289 " 6. Hierarchical Clustering of Genotypes "
290 Symm[Rows=NGeno] Simi
291 Fsim[Simi=Simi] Mean[1...NLoc]; Test=Eucl
292 Hclus[prin=a,d; method=average] Simi ; Amalg=MatSimi ; Permu=PermSimi

**** Average linkage cluster analysis ****

** Merging clusters **

  8    10  95.3
 11    13  95.2
 14    15  94.3
  4     6  93.9
  5     7  93.8
  2     3  91.5
  4     8  91.5
  9    14  90.5
  2     9  89.2
  4     5  89.2
 11    12  85.6
  2     4  84.7
  2    11  81.5
  1     2  74.3

**** Hierarchical clusters ****

** Level      95.0

  8    10

 11    13
** Ungrouped
  1     2   3   9   14   15   4   6   5   7
 12

** Level      90.0

  2     3

  9    14   15

```

```

4      6      8      10
5      7

11     13
** Ungrouped
1      12

** Level      85.0

2      3      9      14     15
4      6      8      10     5      7

11     13     12
** Ungrouped
1

** Level      80.0

2      3      9      14     15     4      6      8      10     5
7      11     13     12
** Ungrouped
1

** Level      75.0

2      3      9      14     15     4      6      8      10     5
7      11     13     12
** Ungrouped
1

** Level      70.0

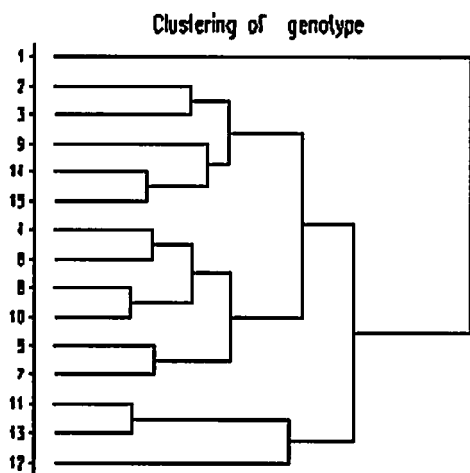
1      2      3      9      14     15     4      6      8      10
5      7      11     13     12

**** Dendrogram ****
** Levels      100.0  90.0  80.0  70.0

1 .....
2 ..... )
3 ..... )
9 ..... )
14 ..... )
15 ..... )
4 ..... )
6 ..... )
8 ..... )
10 ..... )
5 ..... )
7 ..... )
11 ..... )
13 ..... )
12 ..... )

293      DDENDROGRAM [order=given; style=Average]      MatSimi ; permu=PermSimi ; \
294      Title=' Clustering of genotype'
295      Dele[Rede=Y] Simi , MatSimi, PermSimi
296

```



```

297 " 7. Hierarchical Clustering of Environments "
298 " 7. Hierarchical Clustering of Environments "
299
300 Symm[Rows=NLoc] Simi
301 Fsim[Simi=Simi] GMean[1...NGeno]; Test=Eucl
302 Hclus[prin=a,d; method=average] Simi ; Amalg=MatSimi ; Permu=PermSimi

```

**** Average linkage cluster analysis ****

** Merging clusters **

1	6	99.7
3	9	97.5
1	8	97.5
5	10	97.5
2	3	96.9
2	4	96.3
5	7	93.7
2	5	84.5
1	2	57.9

**** Hierarchical clusters ****

** Level 95.0

1	6	8	
2	3	9	4
5	10		

** Ungrouped
7

** Level 90.0

1	6	8	
2	3	9	4

```

5 10 7

** Level      85.0
1 6 8
2 3 9 4
5 10 7

** Level      80.0
1 6 8
2 3 9 4 5 10 7

** Level      75.0
1 6 8
2 3 9 4 5 10 7

** Level      70.0
1 6 8
2 3 9 4 5 10 7

** Level      65.0
1 6 8
2 3 9 4 5 10 7

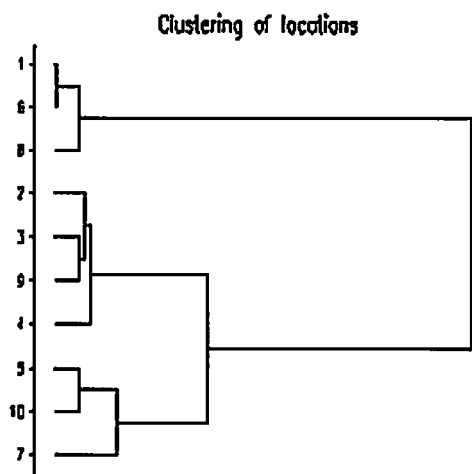
** Level      60.0
1 6 8
2 3 9 4 5 10 7

** Level      55.0
1 6 8 2 3 9 4 5 10 7

**** Dendrogram ****
** Levels 100.0 90.0 80.0 70.0 60.0

1 ..
6 ..)
8 ..).....)
2 ..)
3 ..)
9 ..)
4 ..).....)
5 ..)
10 ..).....)
7 .....).....).....).....)

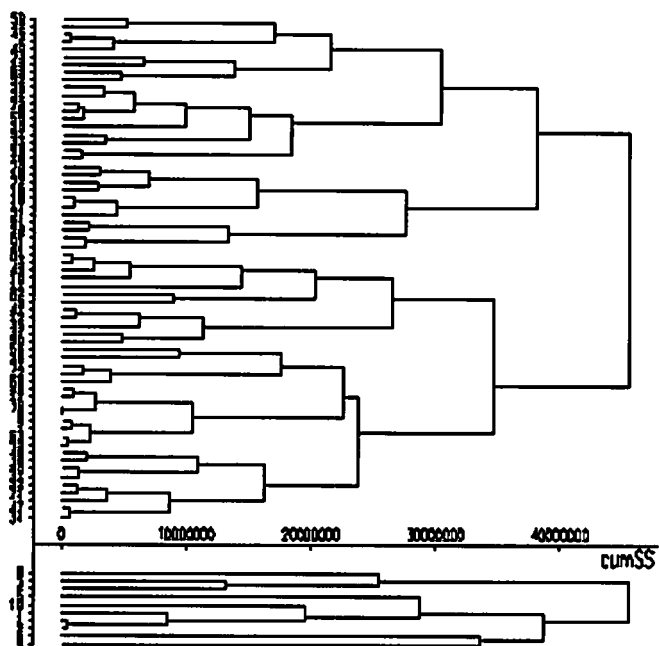
```



```

303
304 DDENDROGRAM [order=given; style=Average]   MatSimi ; permu=PermSimi ; \
305   Title=' Clustering of locations'
306
307 Dele[Rede=Y] Simi , MatSimi, PermSimi
308
309
310 "
-311 8. Clusters genotypes and environments into groups which maximizes
-312   GxE interaction between the groups of genotype and groups of environments
(Corsten and Denis 1990)
-313 8. Clusters genotypes and environments into groups which maximizes
-314   GxE interaction between the groups of genotype and groups of environments
(Corsten and Denis 1990)
-315 "
316
317 "
-318 Scal VarMean : Calc VarMean=PoolMS/NRepAvrg
-319 Scal SStHres : Calc SStHres=(1-GxEI%/100)*GxEISS
-320
-321 Tabu[Class=Genol, Loc1] GEData; Means=TabGE
-322 CINTERACTION [prin=sort,aov,summ,vari,dend ; \
-323   Vari=VarMean; DF=PoolDF; SSTHRES=SStHres] Table=TabGE
-324 "

```



```

325
326
327 " 10. Non-hierarchical clusters of environments "
328 "   Changes are required if more than 3 groups are required"
329
330 If NLoc.lt.NGeno
331   Print '      Number of environments less than number of genotypes ' , \
332         '      resulting in a singular variance-covariance matrix ' , \
333
334         Number of environments less than number of genotypes
335
336         resulting in a singular variance-covariance matrix
337
338   Endif
339
340   Fact[leve=3; nval=NLoc] Grp[3] : & [leve=2; nval=NLoc] Grp[2]
341   Pointer[Valu=GMean[1...NGeno]] EG_data
342   Cluster[prin=c,o; data=EG_data; cri=maha]Ngroups=3,2 ; groups=Grp[3,2]
343
344   For i=1...3
345     Rest LocNum,GMean[1...NGeno] ; Grp[3].eq.i
346     Prin Grp[3],LocNum,GMean[1...NGeno] ;fiel=6
347     Rest LocNum,GMean[1...NGeno]
348   Endf
349
350   For i=1...2
351     Rest LocNum,GMean[1...NGeno] ; Grp[2].eq.i
352     Prin Grp[2],LocNum,GMean[1...NGeno] ;fiel=6
353     Rest LocNum,GMean[1...NGeno]
354   Endf
355
356   Endi
357

```

```

358 " 9. Non-hierarchical clusters of genotypes
-359 9. Non-hierarchical clusters of genotypes
-360 "
361 " Changes are required if more than 3 groups are required "
362
363 If NGeno.lt.NLoc
364 Print '          Number of genotypes less than number of environments' , \
365 '          resulting in a singular variance-covariance matrix'
366 Endif
367
368 If NGeno.ge.NLoc
369 Fact[leve=3; nval=NGeno] Grp[3] : & [leve=2; nvalu=NGeno] Grp[2]
370 Pointer[Valu=Mean[1...NLoc]] GE_data
371 Cluster[prin=c,o; data=GE_data; cri=maha]Ngroups=3,2 ; groups=Grp[3,2]

371.....

**** Non-hierarchical Clustering ****
**** Mahalanobis distance criterion ***

*** Optimum classification ***
*** Number of classes = 3
*** Class contributions not printed ***
*** Criterion value = 8934.36114
*** Classification of units ***
      3      1      2      1      2      2      2      2      2      2      3      2      3
      3      3

*** Optimum classification ***
*** Number of classes = 2
*** Class contributions not printed ***
*** Criterion value = 2444.85203
*** Classification of units ***
      2      2      1      1      2      1      2      1      2      2      2      1      2
      2      2
372
373 For i=1...3
374 Rest GenoNum,Mean[1...NLoc] ; Grp[3].eq.i
375 Prin Grp[3],GenoNum,Mean[1...NLoc] ;fiel=6
376 Rest GenoNum,Mean[1...NLoc]
377 Endf

Grp[3] GenoNum Mean[1] Mean[2] Mean[3] Mean[4] Mean[5] Mean[6] Mean[7] Mean[8]
Mean[9] Mean[10]
1      2      1333      6044      4933      6635      3173      1457      2308      391.7      6300      4933
1      4      1411      5700      5292      6245      4314      699      2785      283.3      6958      3680

Grp[3] GenoNum Mean[1] Mean[2] Mean[3] Mean[4] Mean[5] Mean[6] Mean[7] Mean[8]
Mean[9] Mean[10]
2      3      1341      6556      6100      5967      3423      1050      2611      316.7      5813      4267
2      5      1464      5511      5867      5617      5551      1437      2833      333.3      7558      4013
2      6      1511      5595      6050      5489      3699      1168      3063      275.0      6003      3133
2      7      1411      5933      5350      5737      4987      1451      2740      358.3      5428      4440
2      8      1495      5867      4700      5445      3070      987      3181      375.0      6837      3880
2      9      1516      6389      5133      5700      3897      1899      2596      491.7      5387      4120
2      10     1576      5611      5167      5056      3423      1565      3061      408.3      6200      4280
2      12     1161      6455      6600      5958      5019      1391      3501      483.3      6280      4373

```

```

Grp[3] GenoNum Mean[1] Mean[2] Mean[3] Mean[4] Mean[5] Mean[6] Mean[7] Mean[8]
Mean[9] Mean[10]
3 1 1167 6178 4000 5167 4026 660 2107 441.7 3567 4360
3 11 1511 6189 5400 3833 5135 1254 3048 575.0 5637 4360
3 13 1357 6389 6117 4278 4750 1469 2845 641.7 6360 3893
3 14 1533 6456 4800 4481 3494 1264 2816 366.7 5200 5013
3 15 1354 6167 5750 4259 3333 1203 2817 483.3 4558 4653

```

```

378
379 For i=1...2
380 Rest GenoNum,Mean[1...NLoc] ; Grp[2].eq.i
381 Prin Grp[2],GenoNum,Mean[1...NLoc] ;fiel=6
382 Rest GenoNum,Mean[1...NLoc]
383 Endf

```

```

Grp[2] GenoNum Mean[1] Mean[2] Mean[3] Mean[4] Mean[5] Mean[6] Mean[7] Mean[8]
Mean[9] Mean[10]
1 3 1341 6556 6100 5967 3423 1050 2611 316.7 5813 4267
1 4 1411 5700 5292 6245 4314 699 2785 283.3 6958 3680
1 6 1511 5595 6050 5489 3699 1168 3063 275.0 6003 3133
1 8 1495 5867 4700 5445 3070 987 3181 375.0 6837 3880
1 12 1161 6455 6600 5958 5019 1391 3501 483.3 6280 4373

```

```

Grp[2] GenoNum Mean[1] Mean[2] Mean[3] Mean[4] Mean[5] Mean[6] Mean[7] Mean[8]
Mean[9] Mean[10]
2 1 1167 6178 4000 5167 4026 660 2107 441.7 3567 4360
2 2 1333 6044 4933 6635 3173 1457 2308 391.7 6300 4933
2 5 1464 5511 5867 5617 5551 1437 2833 333.3 7558 4013
2 7 1411 5933 5350 5737 4987 1451 2740 358.3 5428 4440
2 9 1516 6389 5133 5700 3897 1899 2596 491.7 5387 4120
2 10 1576 5611 5167 5056 3423 1565 3061 408.3 6200 4280
2 11 1511 6189 5400 3833 5135 1254 3048 575.0 5637 4360
2 13 1357 6389 6117 4278 4750 1469 2845 641.7 6360 3893
2 14 1533 6456 4800 4481 3494 1264 2816 366.7 5200 5013
2 15 1354 6167 5750 4259 3333 1203 2817 483.3 4558 4653

```

```

384
385 Endif
386
387
388
389 " 11. Principal component analysis for genotypes "
390 " 11. Principal component analysis for genotypes "
391
392 Pointer[values=Mean[1...NLoc]]Data_Loc
393 Matr[rows=NGeno; Colu=NRroots]PCScore
394 Vari[nval=NGeno] PCS[1...NRroots]
395
396 PCP[Print=1,r,t; nroots=NRroots] Data_Loc; Scores=PCScore

```

396.....

**** Principal components analysis ****

```

*** Latent Roots ***
      1          2          3
19652260 11339087 5721755
*** Percentage variation ***
      1          2          3
42.02      24.25      12.24

```

```

*** Trace ***
46763498

```

```

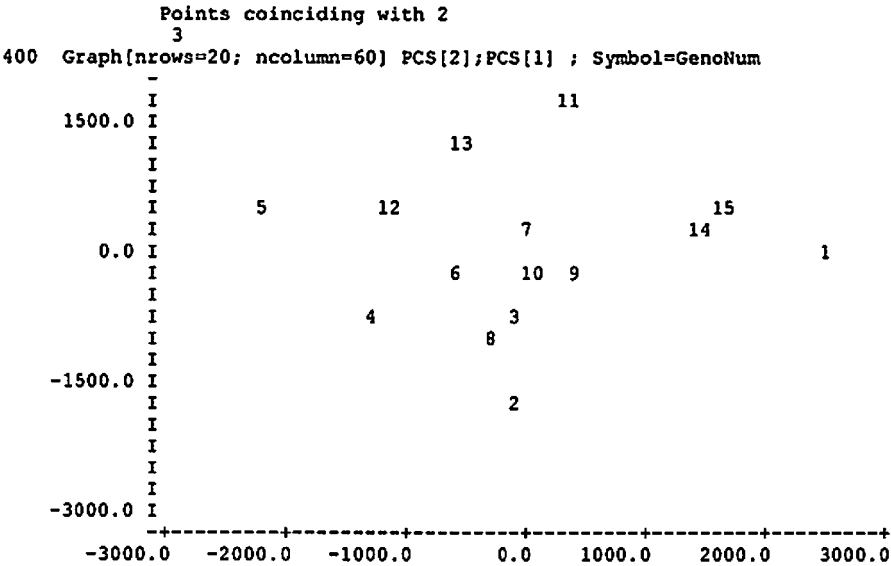
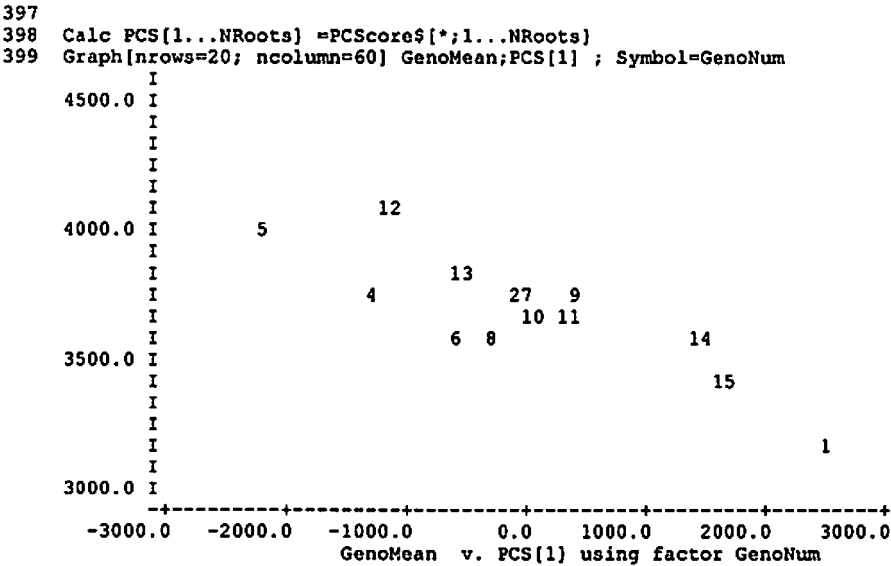
*** Latent Vectors (Loadings) ***
      1          2          3
Mean[1] -0.01595 -0.00287 0.11513
Mean[2] 0.12483 0.08662 -0.06460
Mean[3] -0.34907 0.30239 0.10301
Mean[4] -0.29299 -0.66406 -0.60624
Mean[5] -0.31794 0.63791 -0.63810

```

Mean[6]	-0.05731	0.06721	0.02572
Mean[7]	-0.14832	0.13131	0.24152
Mean[8]	0.02140	0.07354	0.01506
Mean[9]	-0.78072	-0.16108	0.34760
Mean[10]	0.20030	-0.00525	-0.13124

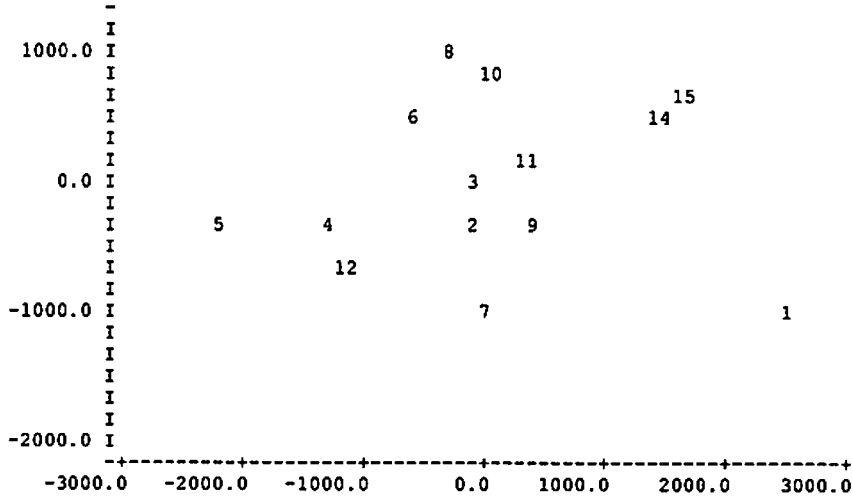
*** Significance tests for equality of final K roots ***

No. (K) Roots	Chi squared	df
2	11.51	2
3	32.02	5
4	37.26	9
5	42.67	14
6	59.62	20
7	73.35	27
8	84.70	35
9	105.01	44
10	129.82	54



PCS[2] v. PCS[1] using factor GenoNum

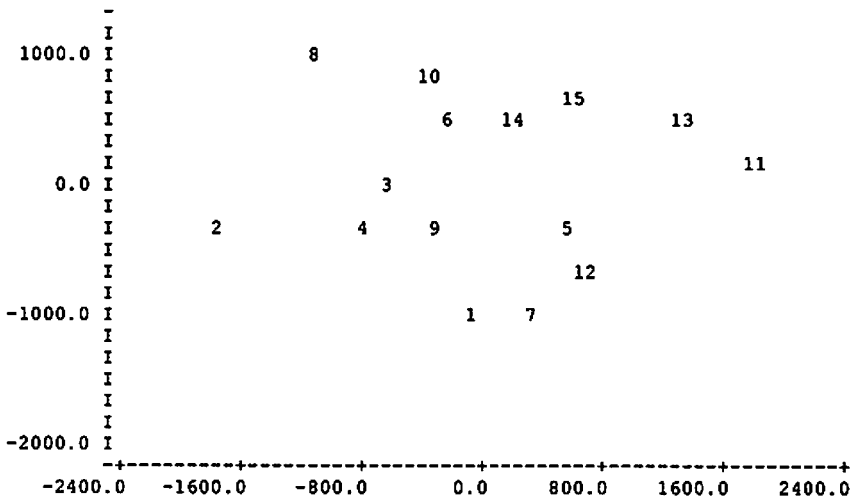
401 Graph[nrows=20; ncolumn=60] PCS[3];PCS[1] ; Symbol=GenoNum



PCS[3] v. PCS[1] using factor GenoNum

Points coinciding with 6
13

402 Graph[nrows=20; ncolumn=60] PCS[3];PCS[2] ; Symbol=GenoNum



PCS[3] v. PCS[2] using factor GenoNum

```

403 Dele Data_Loc
404
405 " 12. Principal component analysis   for environments "
406 " 12. Principal component analysis   for environments "
407 Pointer(values=GMean[1..NGeno])Data_Gen
408 Matr[rows=NLoc; Colu=NRoots]PCScore
409 Vari[nval=NLoc] PCS[1..NRoots]
410
411 PCP[Print=1,r,t; nroots=NRoots] Data_Gen; Scores=PCScore

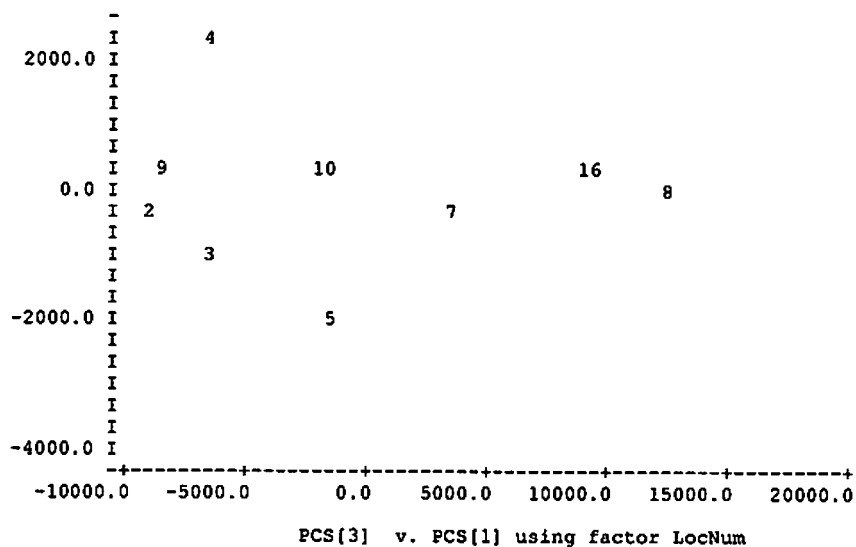
```



```

      I      10
-2000.0 I 2
-----+-----+-----+-----+-----+
-10000.0 -5000.0 0.0 5000.0 10000.0 15000.0 20000.0
      PCS[2] v. PCS[1] using factor LocNum
415 Graph[nrows=20; ncolumn=60] PCS[3];PCS[1] ; Symbol=LocNum

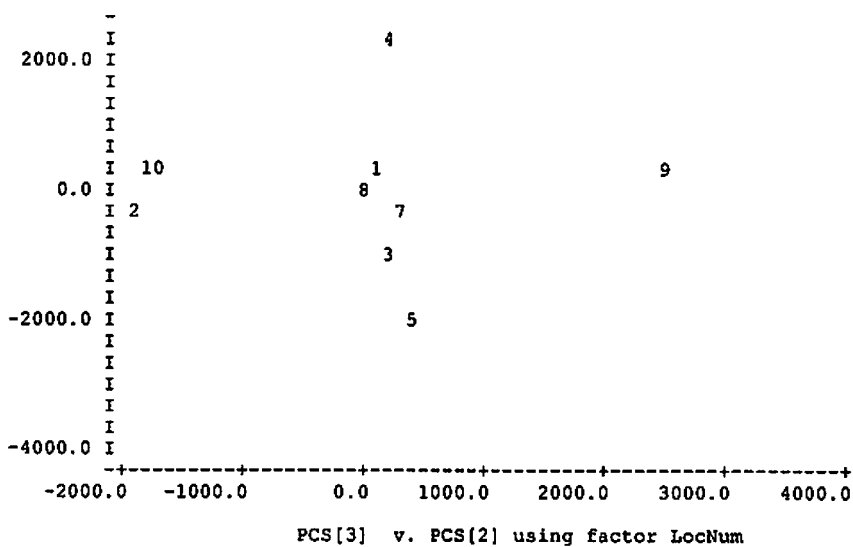
```



```

416 Graph[nrows=20; ncolumn=60] PCS[3];PCS[2] ; Symbol=LocNum

```



Points coinciding with 1
6

```

417
418 " 13. Estimation of variance components and heritabilities"
419 " 13. Estimation of variance components and heritabilities"
420
421
422 " 13.1 From individual environments"

```

```

423
424 Vari[Nvalu=NLoc]RCBSGg2,RCBSGe2,RCBHerit,RCBBias, RCBSch2
425
426 SCAL SGg2,SGe2,h2
427 Scal Vgg,Vge,Vee,Bias, Seh2
428 symm[2] Vcov_r
429
430 For i=1...NLoc
431
432 Rest Yield; Loc.eq.i
433
434 VCOMP[fixed=Rep] RANDOM=Rep+Geno ; cons=pos
435 REML[print=*] Yield
436
437 VKEEP[SIGMA2=SGe2;vcov=Vcov_r] Geno; COMP=SGg2
438 EQUA Vcov_r ; !p(Vgg,Vge,Vee)
439 CALC h2=SGg2/(SGg2+SGe2)
440 CALC One_h22=(1-h2)**2
441 CALC Bias=One_h22*((1-h2)*Vgg-h2*Vge)/(h2*SGe2+SGe2)
442 CALC Seh2=(1-h2)*SQRT(One_h22*Vgg-2*h2*(1-h2)*Vge+Vee*h2**2)/SGe2
443 Calc (RCBSGg2,RCBSGe2,RCBHerit,RCBBias, RCBSch2)$[i]= SGg2,SGe2,h2,Bias,Seh2
444 rest Yield
445 Endf

**** G5W0001 **** Warning (Code CA 7). Statement 10 in For Loop
Command: CALC Seh2=(1-h2)*SQRT(One_h22*Vgg-2*h2*(1-h2)*Vge+Vee*h2**2)/SGe2
Invalid value for argument of function
The first argument of the SQRT function in unit 1 has the value 0.0000

446 Corr[prin=c]LocMean,RCBHerit,RCBSGg2,RCBSGe2

*** Correlation matrix ***
LocMean      1.000
RCBHerit     -0.168      1.000
RCBSGg2       0.483      0.483      1.000
RCBSGe2       0.704     -0.210      0.418      1.000

      LocMean RCBHerit RCBSGg2 RCBSGe2

447 Print LocNum,LocMean,RCBHerit,RCBBias, RCBSch2,RCBSGg2,RCBSGe2 ; fiel=10

LocNum  LocMean  RCBHerit  RCBBias  RCBSch2  RCBSGg2  RCBSGe2
1       1409    0.0000    0.0000    0.0000      0      56612
2       6069    0.0000    0.0000      *         0     433654
3       5417    0.2913    0.1184    0.1731    251218    611261
4       5324    0.0424    0.6085    0.1603    75969    1717369
5       4086    0.3887    0.0924    0.1677    434517    683351
6       1264    0.0000    0.0000    0.0000      0     343770
7       2821    0.3243    0.1100    0.1735    69855    145541
8        415    0.5432    0.0645    0.1471     8587     7220
9       5872    0.3166    0.1105    0.1722    567058    1223885
10      4227    0.0000    0.0000    0.0000      0     892358

448
449 " 13.2 Overall environments"
450
451 SCALAR SGg2,SGe2,SGi2,SGb2,Vgg,Vee,Vii,Vgi,Vge,Vie, h2, Bias,Seh2
452 Symm[3] Vcov_r
453
454 VCOMP[abso=Loc; Fixed=Loc/Rep] RANDOM=Loc+Loc.Rep+Geno+Geno.Loc ; cons=pos
455 REML[print=*] Yield
456
457 VKEEP[SIGMA2=SGe2;vcov=Vcov_r] Geno+Geno.Loc; COMP=SGg2,SGi2
458 EQUA Vcov_r ; !p(Vgg,Vgi,Vii,Vge,Vie,Vee)
459 CALC h2=SGg2/(SGg2+SGi2+SGe2)
460 CALC Bias =h2*(Vgg-h2*(Vgg+Vgi+Vge))/SGg2/SGg2
461 CALC Seh2=Vgg+Vii+Vee+2*(Vgi+Vge+Vie)
462 CALC Seh2=Vgg+h2*h2*Seh2-2*h2*(Vgg+Vgi+Vge)
463 CALC Seh2=h2*SQRT(Seh2)/SGg2
464 Print[iprin=*] 'heritability = ', h2, ' Bias = ', Bias, ' SError= ', Seh2

heritability =      0.02438      Bias =      0.02661      SError=      0.02535

465
466 Clos

```

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